Amygdala volume in schizophrenia: post-mortem study and review of magnetic resonance imaging findings

STEVEN A. CHANCE, MARGARET M. ESIRI and TIMOTHY J. CROW

Background Claims that schizophrenia is a disease of the limbic system have been strengthened by meta-analyses of magnetic resonance imaging (MRI) studies finding reduced hippocampus and amygdala volumes. Some post-mortem studies do not find these abnormalities.

Aims To assess the volume of the amygdala in a series of brains post-mortem.

Method Amygdala volume was estimated using point-counting in both hemispheres of the brains of 10 male and 8 female patients with schizophrenia, and a comparison group of 9 males and 9 females.

Results No significant reduction of amygdala volume was found.

Conclusions Significant volume reduction of the amygdala is not a consistent feature of schizophrenia; findings from early MRI studies using coarse delineation methods may introduce bias to subsequent meta-analyses.

Declaration of interest This work was supported by the Medical Research Council and the SANE Trust.

The limbic system has been suggested to be a possible focus of pathological change in schizophrenia (Torrey & Peterson, 1974; Bogerts, 1997). It is plausible that symptoms of schizophrenia such as inappropriate or flattened affect might relate to a change in the structure of the amygdala and its connections (Benes & Berretta, 2000). Stimulating or damaging the amygdala can induce a range of symptoms (Torrey & Peterson, 1974), and schizophrenia-like psychoses have been associated with epileptic foci in the temporal lobe (Slater & Beard, 1963). There is disagreement between magnetic resonance imaging (MRI) studies and a smaller number of studies in which amygdala volume has been assessed in post-mortem material. A consensus from MRI studies suggests a 5–10% reduction of the amygdala in schizophrenia. However, post-mortem planimetry studies are equivocal. This study aimed to clarify whether the MRI interpretation of amygdala reduction in schizophrenia is apparent in post-mortem tissue, using unbiased methods of estimating volume.

METHOD

Sample Formalin-fixed brain tissue was used post-mortem from subjects who were selected on the basis of the assessment of clinical notes by a consultant psychiatrist (T.J.C. or Dr S. J. Cooper, Belfast). The DSM-IV (American Psychiatric Association, 1994) criteria for schizophrenia and schizoaffective disorder were used in the selection of cases of schizophrenia. Comparison group brains were collected prospectively from individuals who had died without a history of neuropsychiatric disorder, and whose next of kin had given consent for their tissues to be used in medical research. Pathological assessment of tissue samples was carried out by a consultant neuropathologist (M.M.E. or Dr B. McDonald, Oxford), and cases with significant neurological disorders, such as Alzheimer’s disease, were excluded using the criteria of the Consortium to Establish a Registry for Alzheimer’s Disease (Mirra et al, 1991).

Subsequent to sectioning and staining, the series consisted of 36 cases, comprising 10 men and 8 women with schizophrenia, and a comparison group of 9 men and 9 women.

The degree of neuroleptic medication received during life was assessed from case notes, and categorised as ‘little’, ‘average’ or ‘much’. This was based on a clinician’s (T.J.C. or S.A.C.) judgement on the basis of the available clinical records, the variability of information available precluding a more quantitative estimation. A three-category categorisation of causes of death indicates that 14 subjects died of cardiac causes, 9 of respiratory causes and 12 of ‘other’ causes. Of these, 22 deaths were prolonged and 13 were sudden. One subject died of uncertain causes. Other demographic details and potentially confounding variables, including age of symptom onset, duration of illness, age at death, post-mortem interval and fixation time, were noted for statistical analysis (Table 1).

Tissue preparation

Following removal of the leptomeninges and brain-stem, and bisection of the brain, the temporal lobes were removed from the rest of the hemispheres at the posterior end of the sylvian fissure. The temporal lobe was cut into slices 5 mm thick along its antero-posterior axis. Embedding the slices in paraflin resulted in blocks of tissue approximately 3 mm thick owing to shrinkage. Blocks containing the amygdala were exhaustively sectioned from a position posterior to, and random with respect to, the posterior border of the amygdala. For a structure with a regular shape such as the amygdala, a sample of approximately 10 sections should be adequate for the estimation of volume (Gundersen & Jensen, 1987). Selecting sections 25 μm thick with an interslice distance of 650 μm provided about 8–12 sections per amygdala. The sections were counterstained with Luxol fast blue, a myelin stain useful for picking out fibre tracts, and cresyl violet, which stains Nissl’s substance in cell bodies (Fig. 1) (reagents obtained from Raymond A. Lamb Ltd, Eastbourne, UK).

Measurement

A Cavalieri estimate of volume for each amygdala was obtained by stereological
Table I  Demographic variables and details of the brain collection for the amygdala study

<table>
<thead>
<tr>
<th>Variable</th>
<th>Men</th>
<th>Women</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Comparison group (n=9)</td>
<td>Schizophrenia group (n=10)</td>
</tr>
<tr>
<td>Age at onset (years)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mean</td>
<td>NA</td>
<td>28.5</td>
</tr>
<tr>
<td>s.d. (range)</td>
<td>7.7 (21–42)</td>
<td></td>
</tr>
<tr>
<td>Duration of illness (years)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mean</td>
<td>NA</td>
<td>26.5</td>
</tr>
<tr>
<td>s.d. (range)</td>
<td>14.3 (1–44)</td>
<td></td>
</tr>
<tr>
<td>Age at death (years)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mean</td>
<td>68.9</td>
<td>55.0</td>
</tr>
<tr>
<td>s.d. (range)</td>
<td>11.0 (40–80)</td>
<td>13.2 (29–69)</td>
</tr>
<tr>
<td>Post-mortem interval (hours)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mean</td>
<td>42.9</td>
<td>42.0</td>
</tr>
<tr>
<td>s.d. (range)</td>
<td>31.8 (10–98)</td>
<td>34.0 (6–113)</td>
</tr>
<tr>
<td>Fixation period (months)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mean</td>
<td>25.3</td>
<td>46.3</td>
</tr>
<tr>
<td>s.d. (range)</td>
<td>15.9 (1–53)</td>
<td>22.4 (20–91)</td>
</tr>
<tr>
<td>Brain volume (cm³)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mean</td>
<td>1144</td>
<td>1191</td>
</tr>
<tr>
<td>s.d. (range)</td>
<td>115 (995–1290)</td>
<td>104 (994–1378)</td>
</tr>
</tbody>
</table>

1. No accurate information for one case.
NA, not applicable.

Fig. 1  Cross-section of the amygdala and hippocampus. Staining with Luxol fast blue and cresyl violet results in dark staining of areas containing predominantly white matter fibre tracts and lighter staining of nuclei-containing cell bodies. The amygdala lies superior to and overlaps the hippocampus.

Point-counting within the bounds of the amygdala on all sections in which it appeared. A 4-mm² point density grid printed on acetate was laid over each cross-section and viewed under a dissecting microscope × 7 magnification (Fig. 2). Each point represented a volume of 10.8 mm³ (4 mm × 4 mm × 0.675 mm). Three placements of the grid, each time in a random orientation, were used to obtain a mean count for each section. The total count was multiplied by the volume represented by each point for an estimate of total volume. The most anterior section containing the hippocampus and the most anterior section containing the temporal horn were also noted. This enabled a calculation to be made of the percentage of amygdala volume overlapping with each of those structures along the axis of the temporal lobe.

All measures were performed by the same researcher (S.A.C.) who was blind to the diagnosis and gender of the cases. Left, right and mean amygdala volumes were obtained.

Pilot study

A pilot study was used to determine the grid size of points and the number of counts per section required for a satisfactory estimate of amygdala volume. Point counts were made for each of three different grid sizes, varying the number of recounts. Trial grid sizes had interpoint distances of 2 mm, 3 mm and 4 mm, respectively. The 4-mm² grid yielded a total point count of up to about 75 points per amygdala. Three counts...
of this grid gave a coefficient of error of 0.04, which was deemed satisfactory.

Reiability
The volume was estimated twice for 10 randomly selected amygdalas to test intra-rater reliability. This provided an intraclass correlation coefficient of 0.94, suggesting that the reliability of this measure was satisfactory.

Statistics
Statistical analyses were carried out using the software Statistical Package for the Social Sciences, version 9.0 (SPSS, 1998). Unfortunately, for some amygdalas, insufficient tissue was available to provide a systematically random sample of total volume. This was due to an idiosyncratic tissue response to embedding procedures in some tissue blocks which resulted in unsatisfactory sections in a few amygdalas and prohibited their inclusion. Consequently, three cases provided volumetric data only for the right side and four cases only for the left side. Since a repeated-measures analysis of variance (ANOVA) would exclude all seven cases with one side missing, an alternative univariate ANOVA was used to test the mean of both sides for each case, replaced by the value for just one side in cases that did not have both. This approach was used to test for the between-subject effects of diagnosis and gender. A repeated-measures ANOVA was still applied to test the within-subject effects of side using those cases that had data for both sides. In addition, an asymmetry statistic was calculated for these cases: \((R - L)/(R + L) \times 100\), where \(R\) and \(L\) are the values for the right and left sides respectively.

The influence of potentially confounding variables – age at death, post-mortem interval and fixation time – for all cases were examined as variables in separate ANOVAs of groups selected by gender and/or diagnosis. Lifetime medication of the patients with schizophrenia was examined as a factor in ANOVA of the measured variables. Clinical information was felt to be unsatisfactory for correlating symptoms or sub-syndromes with data. All variables, including covariates, fitted with a normal distribution as tested by a Kolmogorov–Smirnov goodness-of-fit test (Zar, 1984) for each gender and diagnosis group. Homogeneity of variance of the measured variables between groups was accepted after testing with a Box’s M test (Glantz & Slinker, 2000) before each ANOVA.

Results
The measurements obtained are summarised in Table 2 and Fig. 3.

Accuracy
The observed coefficients of error for the estimates of volume were less than or equal to 0.1 for both left and right amygdalas in each diagnosis x gender subgroup, indicating that their means offer a satisfactory estimate of the true population means.

Results of analysis of variance
A univariate ANOVA of mean amygdala volume found no significant difference between schizophrenia cases and controls \((F=0.5, \text{ d.f. } 1.30, P=0.47)\), no difference between genders \((F=0.1, \text{ d.f. } 1.30, P=0.73)\) and no interaction between diagnosis and gender. Analysis of covariance (ANCOVA) included fixation time, which was not a significant covariate, and age, which showed a significant negative relationship with volume \((F=5.4, \text{ d.f. } 1.30, P=0.03)\). Inclusion of brain volume in a secondary ANCOVA revealed that it was a significant covariate \((F=10.8, \text{ d.f. } 1.29, P<0.01)\) which had a positive relationship with amygdala volume, and resulted in a reduced \(P\) value in the effect of age \((F=0.9, \text{ d.f. } 1.29, P=0.37)\), suggesting a relationship with age described below. When controlling for brain volume all other results remained, including the lack of difference between patients and controls, and between the genders.

A repeated-measures ANOVA revealed no significant volume asymmetry of the

Table 2  Amygdala measurements in the post-mortem study group. Values are actual measured values uncorrected for tissue shrinkage

<table>
<thead>
<tr>
<th>Variable</th>
<th>Comparison group Mean (s.d.)</th>
<th>Schizophrenia group Mean (s.d.)</th>
<th>Comparison group Mean (s.d.)</th>
<th>Schizophrenia group Mean (s.d.)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Amygdala volume (mm³)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Left</td>
<td>686.5 (165.2) n=8</td>
<td>714.8 (91.9) n=9</td>
<td>628.5 (219.3) n=9</td>
<td>586.6 (138.4) n=7</td>
</tr>
<tr>
<td>Right</td>
<td>607.8 (82.9) n=8</td>
<td>721.2 (14.0) n=9</td>
<td>596.0 (86.5) n=9</td>
<td>545.1 (61.1) n=6</td>
</tr>
<tr>
<td>Mean amygdala volume (mm³)</td>
<td>643.4 (107.5) n=9</td>
<td>720.9 (92.7) n=10</td>
<td>612.3 (182.3) n=9</td>
<td>604.9 (155.0) n=8</td>
</tr>
<tr>
<td>Amygdala symmetry ((R - L)/(R + L) \times 100)</td>
<td>-4.5 (11.4) n=7</td>
<td>-11 (6.5) n=8</td>
<td>-2.2 (14.5) n=9</td>
<td>-3.3 (7.8) n=5</td>
</tr>
<tr>
<td>Percentage of mean amygdala volume overlapping with hippocampus (%)</td>
<td>56.9 (12.5) n=9</td>
<td>60.7 (17.6) n=10</td>
<td>56.0 (18.7) n=9</td>
<td>65.6 (16.5) n=8</td>
</tr>
<tr>
<td>Percentage of mean amygdala volume overlapping with temporal horn (%)</td>
<td>79.6 (11.2) n=9</td>
<td>83.8 (13.4) n=10</td>
<td>71.7 (12.9) n=9</td>
<td>84.2 (12.0) n=8</td>
</tr>
</tbody>
</table>

\(R, L\) are values for right and left sides, respectively.
amygdala and no difference in amygdala symmetry between groups. An ANCOVA found that age and fixation time were not significant covariates (although there was an interaction between fixation time and side: left amygdala smaller with longer fixation time, F=8.8, d.f. 1,32, P<0.01). Brain volume also had no effect on amygdala asymmetry.

The percentage of mean amygdala volume overlapping with the hippocampus, along the temporal lobe axis, was 56% (s.d. 15) in control brains and 63% (s.d. 17) in schizophrenia. The volume overlap with the temporal horn of the ventricles was 76% (s.d. 12) in controls and 84% (s.d. 12) in schizophrenia. No significant correlations were found between amygdala volumes and temporal horn volume.

Artefacts and covariates

As shown in Table 1, the women (F=7.0, d.f. 1,32, P=0.01), particularly those with schizophrenia (F=3.7, d.f. 1,32, P=0.06), were older at death than the men. Consequently, age at death was included as a covariate in all tests of the measured variables. As expected, the total brain volumes were larger in the men than in the women in this series. Amygdala volume was positively correlated with brain volume, which was included as a covariate in secondary ANOVAs of the measured variables to observe the effect of controlling for brain size that might mask changes in the amygdala.

Post-mortem factors

Post-mortem interval did not differ between groups and therefore was not included as a covariate. Fixation time was similarly examined and differed between groups selected by diagnosis (F=25.4, d.f. 1,32, P<0.01) and gender (F=4.0, d.f. 1,32, P=0.06). Consequently, although the distortion due to fixation stabilises after approximately 3 weeks (Quester & Schroder, 1997) and all of the brains studied should therefore have reached a stable state prior to examination, fixation time was included as a covariate in the analysis of amygdala volume.

Clinical factors

Within the group of patients with schizophrenia, age of illness onset and neuroleptic medication were found to have no influence on the mean amygdala volume or the asymmetry of amygdala volumes. The mean level of medication within the schizophrenia group corresponded to a rating of ‘average’ or ‘much’.

DISCUSSION

In a review of MRI studies McCarley et al. (1999) reported that 77% of 30 studies showed a reduction in volume in at least one of the hippocampus, parahippocampal gyrus or amygdala. By meta-analysis Lawrie & Abukmeil (1998) found a median volume reduction of approximately 6% over ten studies measuring the combined amygdala and hippocampus, and about 10% from six studies measuring the amygdala separately. A meta-analysis by Nelson et al. (1998) found a reduction of about 8% from MRI studies measuring the amygdala and hippocampus and about 5% from five studies measuring the amygdala separately. They suggested that inclusion of the amygdala significantly increased the size of the reduction. The most recent meta-analysis, by Wright et al. (2000), found that the amygdala was 94% of its normal size relative to cerebral volume differences in schizophrenia.

In contrast, the post-mortem study by Heckers et al. (1990) found no difference in the volume of the amygdala in 20 patients (see Table 4). Similarly, Pakkenberg’s post-mortem planimetry study (Pakkenberg, 1990) found no difference in volume of the basolateral nucleus, which constitutes a large percentage of the amygdala (the cortico-basolateral nuclei constitute 75% according to Eccles, 1989). These differ from the original finding of reduced amygdala volume in post-mortem material (Bogerts et al., 1985).

Contrary to what might be expected if schizophrenia is a disorder of the limbic system, no volumetric reduction in the amygdala was found in this set of brains. If sustained, this conclusion suggests that the structural changes that have been established in post-mortem studies (Brown et al., 1986; Althshuler et al., 1990), including those of the parahippocampal gyrus in brains from this collection (McDonald et al., 2000), do not extend equally to other parts of the limbic system such as the amygdala and, as previously indicated, may not substantively involve the fornix (Chance et al., 1999).

MRI limitations

Although consistent with other post-mortem findings (Heckers et al., 1990; Pakkenberg, 1990) the absence of a reduction of amygdala size apparently runs counter to the weight of published neuroimaging research. Recent MRI reviews and meta-analyses include studies from the early 1990s, since which time imaging technology has improved. The most obvious limitations of older studies is low scan resolution. Several studies provide only one or two slices through the amygdala (Table 3). Such limited sampling exacerbates the problem of delineating the amygdala, separating it from the hippocampus and the temporal horn of the ventricle. Several authors concluded that the delineation of the amygdala was unreliable and excluded it from their analysis (Flaum et al., 1995), or excluded the region of overlap with the hippocampus (Breier et al., 1992). Many studies (e.g. DeLisi et al., 1991) simply included it with the hippocampus. Some studies (e.g. Bogerts et al., 1990, 1993; Seidman et al., 1999) that divided a segmentation including both amygdala and hippocampus into a posterior portion and an anterior portion (deemed to represent the amygdala) found a volume reduction only when the posterior portion was included.

Table 3 surveys the MRI studies from the period 1988–2000 that have measured the volume of the amygdala in schizophrenia. Studies that provided only one or two slices through the amygdala have been noted. The outcome of the study is reported for the measurement deemed to be closest to an amygdala volume. In some cases this is the anterior part of a hippocampus–amygdala complex. In studies in which the amygdala was defined as a separate entity, the landmark used for the posterior boundary is important. An interslice spacing has been reported for studies that measured more than one slice through the
Amygdala volume in schizophrenia

Table 3  Studies of the amygdala identified by Nelson et al. (1998), Lawrie & Abukmeil (1998) and a Medline search of articles since 1998. The later studies have better scan resolution and assess more than two slices through the amygdala. No slice thickness is given for studies that measure only one slice. For studies with more than two slices through the amygdala, the boundary used to divide the amygdala portion of a segmentation from the hippocampal portion is given. If a subdivision was measured, the findings refer only to the anterior, amygdaloid portion.

<table>
<thead>
<tr>
<th>Study</th>
<th>Number of slices</th>
<th>Interslice spacing (mm)</th>
<th>Arbitrary boundary</th>
<th>Significant volume reduction?</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>&gt; 3 mm</td>
<td>≤ 3 mm</td>
<td></td>
</tr>
<tr>
<td>Kelsoe et al (1988)</td>
<td>2</td>
<td>10</td>
<td></td>
<td>No</td>
</tr>
<tr>
<td>Suddath et al (1989)</td>
<td>1</td>
<td></td>
<td></td>
<td>Yes</td>
</tr>
<tr>
<td>Barta et al (1990)</td>
<td>2</td>
<td>3</td>
<td></td>
<td>Yes</td>
</tr>
<tr>
<td>Bogerts et al (1990)</td>
<td>&gt; 2</td>
<td>3.1</td>
<td>Mamillary bodies</td>
<td>No</td>
</tr>
<tr>
<td>Dauphinais et al (1990)</td>
<td>2</td>
<td>10</td>
<td></td>
<td>Yes</td>
</tr>
<tr>
<td>Suddath et al (1990)</td>
<td>1</td>
<td></td>
<td></td>
<td>No</td>
</tr>
<tr>
<td>Blackwood et al (1991)</td>
<td>&gt; 2</td>
<td>8</td>
<td>Undefined</td>
<td>No</td>
</tr>
<tr>
<td>DeLisi et al (1991)</td>
<td>&gt; 2</td>
<td>7</td>
<td>Combined measure</td>
<td>No</td>
</tr>
<tr>
<td>Breier et al (1992)</td>
<td>&gt; 2</td>
<td>3</td>
<td>Overlap excluded</td>
<td>Yes</td>
</tr>
<tr>
<td>Hoff et al (1992)</td>
<td>&gt; 2</td>
<td>7</td>
<td>Pons</td>
<td>No</td>
</tr>
<tr>
<td>Shenton et al (1992)</td>
<td>&gt; 2</td>
<td>1.5</td>
<td>Mamillary bodies</td>
<td>Yes</td>
</tr>
<tr>
<td>Swaze et al (1992)</td>
<td>2</td>
<td>10</td>
<td></td>
<td>No</td>
</tr>
<tr>
<td>Bogerts et al (1993)</td>
<td>&gt; 2</td>
<td>3.1</td>
<td>Mamillary bodies</td>
<td>No</td>
</tr>
<tr>
<td>Marsh et al (1994)</td>
<td>1</td>
<td></td>
<td></td>
<td>Yes</td>
</tr>
<tr>
<td>Rossi et al (1994)</td>
<td>2</td>
<td>6</td>
<td></td>
<td>Yes</td>
</tr>
<tr>
<td>Becker et al (1996)</td>
<td>&gt; 2</td>
<td>4</td>
<td>Mamillary bodies</td>
<td>No</td>
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<tr>
<td>Pearson et al (1997)</td>
<td>2</td>
<td>3</td>
<td></td>
<td>Yes</td>
</tr>
<tr>
<td>Hirayasu et al (1998)</td>
<td>&gt; 2</td>
<td>1.5</td>
<td>Mamillary bodies</td>
<td>No</td>
</tr>
<tr>
<td>Giedd et al (1999)</td>
<td>&gt; 2</td>
<td>2</td>
<td>Tracing</td>
<td>No</td>
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<tr>
<td>Sanderson et al (1999)</td>
<td>&gt; 2</td>
<td>1.9</td>
<td>Combined measure</td>
<td>Yes</td>
</tr>
<tr>
<td>Seidman et al (1999)</td>
<td>&gt; 2</td>
<td>3</td>
<td>10/24 of AC–PC distance (automated)</td>
<td>No</td>
</tr>
<tr>
<td>Altschuler et al (2000)</td>
<td>&gt; 2</td>
<td>1.4</td>
<td>Tracing</td>
<td>No</td>
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<tr>
<td>Niemann et al (2000)</td>
<td>&gt; 2</td>
<td>2.2</td>
<td>Tracing</td>
<td>No</td>
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<tr>
<td>Staal et al (2000)</td>
<td>&gt; 2</td>
<td>1.6</td>
<td>Optic tract</td>
<td>No</td>
</tr>
</tbody>
</table>

AC, anterior commissure; PC, posterior commissure.

Table 4  Post-mortem studies of the amygdala in schizophrenia

<table>
<thead>
<tr>
<th>Study</th>
<th>Structure</th>
<th>Volume (cm³)</th>
<th>Significant difference?</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>Control mean (s.d.)</td>
<td>Schizophrenia mean (s.d.)</td>
</tr>
<tr>
<td>Bogerts et al (1985)</td>
<td>Amygdala</td>
<td>1.62 (0.78) n=8</td>
<td>1.27 (0.78) n=9</td>
</tr>
<tr>
<td>Heckers et al (1990)</td>
<td>Amygdala</td>
<td>1.20 (0.14) n=20</td>
<td>1.29 (0.22) n=20</td>
</tr>
<tr>
<td>Pakkenberg (1990)</td>
<td>Basolateral amygdala</td>
<td>0.19² n=5</td>
<td>0.19² n=5</td>
</tr>
</tbody>
</table>

1. NS denotes 'not significant' where a p value was not given in the paper.
2. Approximate values derived from scatter plot (also appear to be uncorrected for tissue shrinkage).

Limitations of meta-analysis

Although apparently consistent evidence of amygdala volume reduction emerges from compilations of studies, such meta-analyses conceal a potential source of bias. In most studies a landmark external to the temporal lobe was used either to mark the division into anterior and posterior portions, or to begin segmentation of the amygdala. Few MRI studies that provide more than two slices through the amygdala delineate it without the use of an external landmark (Fig. 4). The landmarks include the mamillary bodies (Bogerts et al, 1990, 1993; Shenton et al, 1992; Rossi et al, 1994; Becker et al, 1996; Hirayasu et al, 1998), the optic tract (Staal et al, 2000), the pons (Holf et al, 1992), and the anterior and posterior commissures (Seidman et al, 1996).
However, the temporal lobes have been reported to be preferentially shortened in schizophrenia (Bartzokis et al., 1996). For example, in the present series of brains, measurement from the pole to the posterior sylvian fissure (Highley et al., 1998), and to the ventricular trigone (further details available from the author upon request), have found shorter temporal lobes despite controlling for brain size. In this situation, landmarks that lie outside the temporal lobe in cases of schizophrenia will be further forwards relative to structures within the temporal lobe. Consequently, the use of external landmarks in MRI studies could constitute one systematic source of error, yielding smaller estimates of amygdala volume such as are identified in the meta-analyses.

Furthermore, this study found that at least half of the volume of the amygdala overlaps the hippocampus along the axis of the temporal lobe. This emphasises the necessity of clearly differentiating the two structures to obtain accurate volumes. Although the mean amygdala volume was not changed in this study, the percentage of its volume overlapped by the neighbouring hippocampus and temporal horn was slightly more in the brains from subjects who had schizophrenia. The greater surface of amygdala bordered by the high contrast of an enlarged temporal horn in schizophrenia could lead to a more conservative segmentation of the amygdala on MRI scans.

**Hemisphere and gender**

No significant asymmetry of amygdala volume was found in these brains – a finding similar to that of Heckers et al. (1990). No difference of asymmetry was noted between patients with schizophrenia and comparison subjects. In contrast to the finding of Fukuzako et al. (1997), who showed that a relatively larger right hippocampus was correlated with later age of illness onset, no correlation between amygdala asymmetry and age of onset was found. This was accompanied by an absence of significant difference between the genders, similar to the findings of Bryant et al. (1999), although the mean amygdala volume in the women seemed to be a little less than that in the men (see Table 2) – presumably associated with the normal human sexual dimorphism of total brain size.

**Interpreting normal volume**

The negative findings of this study (particularly in relation to imaging studies) raise the possibility of a type II error (i.e. a failure to identify a difference between groups because of small sample size). A power analysis using the Statmate program, version 1.01 (GraphPad Software, San Diego, USA) indicates that this study had about 15% power to detect a change of 5–7% of amygdala volume, as suggested by MRI meta-analyses. This rises to about 30% power to detect a change of 10% volume, and 80% power to detect a change of 20% (power = 1 – p (type II error); the model was a two-samples t-test with α = 0.05). The observed coefficients of error, which take into account sample size to determine how good an estimate of the true population mean is provided by the sample mean, were all satisfactory, at less than or equal to 0.1.

The absence of a gross volumetric reduction in this study does not discount a cytoarchitectural or neurochemical disturbance. Increased dopamine innervations on the left (Reynolds, 1983) and reduced binding of γ-amino butyric acid (Simpson et al., 1989) in the amygdala in schizophrenia, have been interpreted as a loss of inhibition which might induce positive symptoms (Reynolds, 1995). Such changes may reflect a change in connectivity, for example of the dopaminergic afferent fibres, rather than an alteration in gross volume.

The study was limited by the use of elderly, medicated subjects, and Bogerts et al. (1991) have suggested that age-related brain atrophy may obscure reductions in limbic structures. In our study the negative association of age with volume could act to reduce any apparent diminution in volume in the subjects with schizophrenia, since the control subjects were on average 5 years older (means: controls 69 years, patients 64 years). However, age was controlled for as a covariate in the analyses of volume. The use of MRI can avoid these complications.

Magnetic resonance imaging is most appropriate for assessment of macroscopic measurements, while post-mortem examination is still the only option at a microscopic scale. Currently, the amygdala stands at the limit of structures that can be satisfactorily determined using MRI. Recent and improving methods of assessment, which make use of visual tracing of the amygdala–hippocampus boundary, using the alveus and local anatomical features rather than other external landmarks, should provide a more reliable source of measurements (Kates et al., 1997; Niemann et al., 2000; Pruessner et al., 2000). Further studies are required to clarify which, if any, components of the limbic system are affected in schizophrenia. The current study supports the conclusion that volume reductions of the amygdala in schizophrenia are not large.
and that small reductions reported in MRI may be due to coarse delineation methods that could introduce bias to subsequent meta-analyses.

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Amygdala volume in schizophrenia: post-mortem study and review of magnetic resonance imaging findings
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