Cognitive dysfunction in first-episode psychosis: the processing speed hypothesis*

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Background  Speed of processing is a cognitive process underlying cognitive dysfunction in people with chronic schizophrenia.

Aims  To investigate the contribution of speed of processing to the cognitive deficits observed in a representative large sample with first-episode schizophrenia.

Method  People with a diagnosis of first-episode schizophrenia-spectrum disorders (n=26) and healthy controls (n=28) were compared on several cognitive measures before and after controlling for speed of processing.

Results  Before controlling for speed of processing, patients and controls differed significantly on all cognitive measures. All significant differences in cognitive functioning disappeared when the result of the Digital Symbol Substitution Test was included as an additional covariate.

Conclusions  Speed of information processing may be considered a core cognitive deficit in schizophrenia and might be mediating a broader diversity of cognitive disturbances.

Declaration of interest  None. Funding detailed in Acknowledgements.

Participants

From February 2001 to October 2005 all referrals to PAFIP were screened for the following eligibility criteria: (a) aged 15–60 years; (b) meeting DSM–IV criteria (American Psychiatric Association, 1994) for a principal diagnosis of schizophrenia, schizophreniform disorder, schizoaffective disorder, brief reactive psychosis, schizotypal personality disorder or psychosis not otherwise specified; (c) no prior treatment with antipsychotic medication or, if previously treated, a total lifetime of adequate antipsychotic treatment of less than 6 weeks. Patients meeting these criteria and their families provided written informed consent for inclusion in the PAFIP programme. The diagnoses were confirmed using the Structured Clinical Interview for DSM–IV (SCID–I; First et al, 2001) by an expert psychiatrist 6 months after the initial contact. Only those people with mental retardation (DSM–IV criteria) or drug dependence (DSM–IV criteria) were excluded.

Out of 174 consecutive admissions that met the criteria for enrolment, 43 people declined to participate in the study. A final sample of 131 patients (85 males, 46 females) completed the baseline neurocognitive assessment. No significant differences between patients who did and did not complete cognitive assessment were found in relevant variables such as age, gender, duration of illness, or clinical severity. From these 131 patients, only 126 with a final diagnosis of schizophrenia, schizophreniform disorder, brief psychotic disorder and psychosis not otherwise specified at 6 months by SCID–I were considered for this study.

At the time of the cognitive evaluation 42 patients were taking conventional (haloperidol) and 84 atypical antipsychotics (41 olanzapine and 43 risperidone). Clinical symptoms were rated using the Scale for the Assessment of Negative Symptoms (SANS; Andreasen, 1983) and the Scale for the Assessment of Positive Symptoms (SAPS; Andreasen, 1984). Duration of untreated illness was defined as the time from the first non-specific symptoms related to psychosis (for such a symptom to be considered, there should be no return to the previous stable level of functioning) to the initiation of adequate antipsychotic drug treatment. Duration of untreated psychosis was defined as the time from the first continuous (present most of the time) psychotic

symptom to initiation of adequate anti-psychotic drug treatment.

A sample of healthy controls ($n=28$) recruited from the local area through advertisement was submitted to the same cognitive assessment. Antecedents of psychiatric disturbance in the control group were excluded by means of the Comprehensive Assessment of Symptoms and History (CASH; Andreasen et al, 1992).

**Neuropsychological assessment**

Patients and controls were submitted to a broad neuropsychological battery. This was carried out once clinical stability had been achieved in order to maximise collaboration. This occurred typically at a mean of 10.72 (s.d.=3.97) weeks after intake into the programme.

In order to reduce type I error on statistical analyses, variables obtained from neuropsychological tests were grouped into cognitive domains following results from previous factor analysis. A detailed description of the cognitive battery employed, the procedure for factor analysis and the final composition of cognitive factors is provided elsewhere (Gonzalez-Blanch et al, 2007).

From eight basic cognitive domains, we decided to examine those key cognitive dimensions with a more pronounced impairment (around 1 s.d. below healthy controls): executive functions, verbal memory, attention and motor dexterity were considered as dependent variables.

The Digit Symbol Substitution Test from the WAIS–III battery (Weschler, 1999) was employed as a measure of speed of processing (Salthouse, 1992; Brebion et al, 1998) to be covaried on statistical analyses. This test was included in the executive functions dimension on previous factor analysis. Since a variable should not be used as a covariate and a dependent variable at the same time, we chose to examine tests related to executive function independently. Span of Backward Digits Test (Weschler, 1999), Trail-Making Test Part B (TMT–B; Lezak, 1995) and a fluency test (FAS; Lezak, 1995) were used as representative of working memory and executive functions.

Verbal memory was assessed using the Rey Auditory Verbal Learning Test (RAVLT; Rey, 1964) immediate memory span, total learning, recall following short and long delay periods, and recognition. Motor dexterity was assessed using the grooved pegboard (both hands) and sustained attention/vigilance the Continuous Performance Test–Degraded Stimulus (CPT–DS; Cegalis & Bowlin, 1991) hits and reaction time and Brief Test of Attention (BTA; Schretlen et al, 1996). The vocabulary test from WAIS–III was used to control for premorbid IQ (Lezak, 1995).

**Statistical analysis**

The independent samples $t$-test was used to compare patients and controls on age, years of education and premorbid IQ. The $\chi^2$ test was used to compare gender distribution on both samples. Univariate analysis of covariance (ANCOVA) was used to compare both groups on each of the dependent variables separately. Univariate instead of multivariate ANCOVA was used because of missing data. This caveat precluded the use of multivariate testing since it involved an important loss of sample size. However, we considered the number of dependent variables small enough that errors derived from multiple comparisons would not be significant. Two sets of ANCOVA were carried out. On the first only those socio-demographic variables that differed between patients and controls were covaried. On the second results of the Digit Symbol Substitution Test (speed of processing) were also included as a covariate.

**RESULTS**

**Sample**

The socio-demographic and clinical characteristics of patients and controls are presented in Table 1. The two groups did not differ on age ($t=0.63$, $d.f.=148$, $P=0.53$) or gender distribution ($\chi^2=3.07$, $P=0.08$), although there was a trend-level difference. Patients and healthy controls differed in terms of years of education ($t=-2.15$, $d.f.=148$, $P=0.03$) and premorbid IQ ($t=-2.27$, $d.f.=148$, $P=0.03$). Therefore on the subsequent analysis, gender, years of education and premorbid IQ were always included as covariates.

There were no differences in socio-demographic and cognitive characteristics when patients with schizophrenia and with other psychoses were compared (data not shown).

<table>
<thead>
<tr>
<th>Characteristics</th>
<th>Patients ($n=126$)</th>
<th>Healthy controls ($n=28$)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age, years: mean (s.d.)</td>
<td>26.9 (7.3)</td>
<td>25.67 (6.3)</td>
</tr>
<tr>
<td>Education, years: mean (s.d.)</td>
<td>10.32 (3.3)</td>
<td>11.74 (2.3)</td>
</tr>
<tr>
<td>Premorbid IQ: mean (s.d.)</td>
<td>9.06 (3.1)</td>
<td>10.44 (2.5)</td>
</tr>
<tr>
<td>DU, months: mean (s.d.)</td>
<td>24.63 (33.1)</td>
<td></td>
</tr>
<tr>
<td>DUP, months: mean (s.d.)</td>
<td>12.13 (24.3)</td>
<td></td>
</tr>
<tr>
<td>SAPS total score at intake</td>
<td>12.78 (41)</td>
<td></td>
</tr>
<tr>
<td>SANS total score at intake</td>
<td>7.79 (65)</td>
<td></td>
</tr>
<tr>
<td>Males, n (%)</td>
<td>81 (64.3)</td>
<td>13 (46.4)</td>
</tr>
<tr>
<td>Hospitalisation required, n (%)</td>
<td>82 (65.1)</td>
<td></td>
</tr>
<tr>
<td>Family history of psychosis, n (%)</td>
<td>23 (18.3)</td>
<td></td>
</tr>
<tr>
<td>Normal employment, n (%)</td>
<td>37 (29.4)</td>
<td></td>
</tr>
<tr>
<td>Diagnosis, n (%)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Schizophrenia</td>
<td>77 (61.1)</td>
<td></td>
</tr>
<tr>
<td>Schizophreniform disorder</td>
<td>34 (27.0)</td>
<td></td>
</tr>
<tr>
<td>Brief psychotic disorder</td>
<td>8 (6.3)</td>
<td></td>
</tr>
<tr>
<td>Psychosis not otherwise specified</td>
<td>7 (5.6)</td>
<td></td>
</tr>
<tr>
<td>Pharmacological treatment, n (%)</td>
<td>42 (33.3)</td>
<td></td>
</tr>
<tr>
<td>Haloperidol</td>
<td>41 (32.5)</td>
<td></td>
</tr>
<tr>
<td>Olanzapine</td>
<td>43 (34.1)</td>
<td></td>
</tr>
</tbody>
</table>

DU, duration of untreated illness; DUP, duration of untreated psychosis.
I. Total scores were calculated using sums of global scores of each symptom.
Cognitive performance

Patients and controls differed greatly on performance on the Digit Symbol Substitution Test ($t = -7.51$, d.f. = 148, $P < 0.001$). This significant difference remained after controlling for gender, years of education and premorbid IQ (ANCOVA: $F = 52.40$, $P < 0.001$) and for all socio-demographic and cognitive variables together (ANCOVA: $F = 16.552$, $P < 0.001$).

Differences in cognitive performance between patients and controls remained when gender, education and premorbid IQ were included as covariates. After inclusion of the speed of processing as an additional covariate, none of the differences reached the established significance. Table 2 shows the results of ANCOVA before and after controlling for speed of processing.

**DISCUSSION**

The current study was driven by the hypothesis that deficits in cognition in schizophrenia may be determined by a slower speed of processing of cognitive performance. We compared the cognitive performance of people with schizophrenia and controls before and after controlling for the effect of speed of information processing on cognitive functioning. Interestingly, the results showed that when the influence of speed of processing was removed the cognitive deficits observed in people with first-episode schizophrenia disappeared. Moreover, the significant differences between patients and healthy volunteers in the performance of the Digit Symbol Substitution Test reveal that speed of processing is severely impaired in the early phases of psychosis. Taken together these results suggest that speed of information processing may be a core cognitive deficit in schizophrenia and might mediate a broader diversity of cognitive disturbances.

Our results confirm and extend previous findings in chronic schizophrenia on the influence of speed of information processing on memory and working memory (Brebinon et al., 1998, 2000; Hartman et al., 2002). We have also found that a more diverse range of cognitive functions, including attention, motor dexterity and executive functions, are also influenced by speed of processing, and that this is already evident during the early phases of the disease. It has been proposed that age-related cognitive decline might be associated with a decrease in the speed with which many processing operations can be executed and that this reduction leads to impairments in cognitive functioning (Salthehouse, 1996).

Andreasen and colleagues (1998) proposed the cognitive dysmetria model in schizophrenia. This posits that a neural misconnection in cortical-cerebellar-thalamic-cortical circuitry in people with schizophrenia leads to an impairment in synchrony or the smooth coordination of mental processes. When synchrony is impaired, there is cognitive dysmetria, and the impairment in this basic cognitive process defines the phenotype of schizophrenia and produces its diversity of symptoms. We hypothesise that processing speed is one ‘candidate cognitive dysfunction’ that could explain the myriad symptoms of schizophrenia.

Neuroimaging studies in multiple sclerosis (Sanfilipo et al., 2006), stroke (Jokinen et al., 2005) and normal ageing (Gunning-Dixon & Raz, 2000) have demonstrated that deficits in speed of processing are associated with abnormalities in cerebral white matter. It is noteworthy that abnormalities in white matter have also been related to deficits in processing speed in schizophrenia (Antonova et al., 2003).

In accordance with the cognitive dysmetria hypothesis, a deficit in speed of processing in schizophrenia might reflect the disruption of neural activation transmission owing to white matter alteration. This might secondarily affect higher-order processing, leading to functional disturbances. Cognitive impairment in schizophrenia has great relevance for the evolution of the illness, especially the level of general functionality and autonomy that patients eventually achieve (Green, 1996). This emphasises the potential value of addressing cognitive dysfunction as a primary target of therapeutic intervention. Cognitive remediation in schizophrenia has already proved to be valuable in terms of amelioration of problems with social functioning (Wykes et al., 2003). If our hypothesis proves to be true, pharmacological or neurorehabilitation-based therapeutic approaches aimed at remediating specific cognitive deficits (e.g. memory or executive functions) should take into account the effects of an impaired speed of information processing which might may be understood as a core cognitive deficit.

One of the limitations of this study might be the validity of the test used to measure speed of processing. Although the Digit Symbol Substitution Test has been traditionally used as a measure of processing speed (Salthehouse, 1992, 1993; Lindenberger et al., 1993; Brebinon et al., 1998), it might also be measuring a large
number of other cognitive processes. This limitation applies to most of the neuropsychological tasks and emphasizes the need for more specific cognitive tasks for research purposes. A second limitation of this study is that patients had been treated prior to cognitive evaluation. Although treatment was minimal, antipsychotics might have produced changes in cognition. If medication had some particular significant effect on speed of information processing this might have biased the interpretation of our results. However, to our knowledge there are no reports of a specific effect of antipsychotics on speed of processing. Therefore, the significant relationship between cognitive functioning and processing speed is unlikely to be due to medication.

In conclusion, we suggest that speed of processing might be a core cognitive function in schizophrenia. Thus, cognitive deficits described in people with schizophrenia from the very early stages of the illness might be determined by a slower speed of information processing. Our findings provide further support for recent theories of a neural basis of schizophrenia, such as the cognitive dysmetria theory, and suggest new strategies for neurorehabilitation research.

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