Antipsychotics and borderline personality disorder

I congratulate Lieb et al on their excellent systematic review. However, it is interesting that studies until June 2008 were included in this review; moreover, that in January 2009 the National Institute for Health and Clinical Excellence (NICE) guidelines advised that ‘drug treatment should not be used specifically for borderline personality disorder or for the individual symptoms or behaviour associated with the disorder’.

I am surprised that there were no randomised controlled trials (RCTs) available at the time of study on the usefulness of quetiapine, although some RCTs of aripiprazole and olanzapine were. A few open-label studies have been done highlighting the usefulness of quetiapine in reducing impulsivity and affective symptoms, and it is evident in clinical practice that it does have some beneficial effects on mood instability and aggression.

It is a pity that forest plotting could not be done, which would have shown how much variation existed among studies and the degree of precision of each study, although one can understand the various difficulties faced by the authors.

Lastly, I would like to seek clarification regarding somewhat conflicting statements in the paragraph ‘Implications for practice and research’; it initially states ‘nor can low-dose antipsychotics be advised for cognitive–perceptual symptoms as earlier recommended by the American Psychiatric Association Practice Guidelines’, but later states ‘the SGAs (aripiprazole, olanzapine) should be the first choice for treating cognitive–perceptual symptoms’.


Authors’ reply: We agree that the conclusions from NICE and our review are surprisingly different, considering similar literature search periods and widely similar inclusion criteria for primary studies. However, our scope was to assess and evaluate ‘the mere evidence’ of clinical outcomes. The National Institute for Health and Clinical Excellence, in contrast, aims at the formulation of instructional recommendations for the British National Health Service. Thus, the steering group may have had a broader view and considered additional criteria such as cost-effectiveness within a complex system of healthcare. Differences may therefore, at least in part, stem from different perspectives and scopes: the assessment of the mere evidence and the formulation of instructional guidelines.

Indeed, there were and still are no RCTs on quetiapine available. We are aware of one RCT (the Verkes Borderline Study) that has not been published (yet). Thank you for the reference list. There are two more open-label trials of quetiapine in borderline personality disorder. However, this list is not necessarily exhaustive.

We agree that forest plotting would have contributed to a more immediate understanding of the evidence. However, may we refer you to the full Cochrane review which is to be published soon in the Cochrane Library. Forest plots will be provided there whenever appropriate.

Finally, we thank you for indicating this passage which is indeed liable to misunderstanding. The American Psychiatric Association guidelines recommend low-dose antipsychotics in general, whereas our findings indicate that second-generation antipsychotics are supported by the current RCT evidence in particular. This development of a shift towards second-generation antipsychotics has been foreshadowed by John M. Oldham in his guideline watch of 2005 but, to our knowledge, the original guideline recommendations have not been modified since.


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Sample bias may obscure results

Di Forti et al present their paper as further evidence of the link between high-potency cannabis and psychosis. Obviously, a major issue in case–control studies is the sampling, and any difference between case and control groups needs to be carefully considered. The authors state that ‘there was no significant difference between the cases and control groups in age, gender, ethnicity, educational qualifications or employment status at time of assessment’. However, I would raise concerns about the employment status
of the participants and I would respectfully highlight that this statement does not seem consistent with the information provided in the accompanying table of sample characteristics. This table states that 38.4% of cases and 43.2% of controls were unemployed. The percentages in this table have some inaccurate rounding but more worryingly, contrary to the authors’ report, there is a clear statistically significant difference ($P = 0.001$ using a z-test for proportions).

This also seems to be a highly relevant and clinically significant difference that may have introduced considerable bias into this study and merited the attention of the 14 authors. In the discussion the authors state ‘the increased availability of skunk cannot alone explain why our control group members are less likely to prefer higher-potency types than the cases group across time’. The requirement to hold down a job may be a highly significant reason why controls smoked cannabis of lesser potency less often than the unemployed. Moreover, individuals who are unemployed are highly likely to have poorer social and health status, which further serves to obscure the true role of cannabis in this study.

Authors’ reply: Among the sociodemographic variables we reported in Table 1, it is correct to point out that unemployment rates are statistically significantly higher in the cases compared with controls ($P<0.001$). This difference has already been reported in previous epidemiological studies and there is no evidence that this arises from a bias in the sample selection. However, it is rather a potential confounder. In our paper we did not discuss if or how employment status might have influenced our findings, because, together with other relevant variables, we controlled for it in the statistical analyses. Thus, the higher rate of unemployment in cases than controls might partially account for the drop of the crude odds ratio (OR) of 8.1 (95% CI 4.6–13.5) to the adjusted one (OR=6.8, 95% CI 2.6–25.4), which occurred when we controlled for confounders including unemployment. However, the odds ratio still remains strikingly high and statistically significant ($P<0.05$), indicating that our findings cannot be explained by the effect of employment status or by any of the other social variables listed.

Lastly, we wish to comment on the suggestion that controls’ preference for low-potency cannabis might be consequent to their need to continue being able to work. Would this not indicate that high-potency cannabis is more likely to negatively affect social functioning perhaps via its detrimental effect on mental health? Exactly what our findings suggest.


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Corrections

Superior temporal gyrus volume in antipsychotic-naive people at risk of psychosis. *BJP*, 196, 206–211. The second sentence of the Method (p. 206) should read: Those recruited were aged 14–30 years, had not experienced a previous psychotic episode, had never received any psychotropic medication (antipsychotics, antidepressants, mood stabilisers or benzodiazepines) and had an IQ score above 70, assessed with the National Adult Reading Test.

Bringing new life into psychiatry – extra. *BJP*, 196, 248. The doi was printed incorrectly and should be: 10.1192/bjp.196.3.248a. The online version has been corrected in deviation from print and in accordance with this correction.

Recent trends in the incidence of recorded depression in primary care. *BJP*, 195, 520–524. In the key to Fig. 1 (p. 522) ‘Depression’ and ‘Combined’ are transposed. The correct figure is reproduced below.

![Fig. 1](image-url) Incidence of diagnosed depression and depressive symptoms.

PYAR, person-years at risk.
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