Conduct problems are common in youth, are potent risks for poor psychiatric and psychosocial outcomes in adulthood, and carry high costs to society.1,2 An important outstanding question concerns the extent to which children with conduct problems are at increased risk for adult depression.3–5 Others have failed to find an association.6 Confirming the extent and nature of the link between conduct problems and early adult depression is, however, important from a clinical and public health point of view. There are effective treatments for conduct problems.7 If there were a causal relationship between depression and conduct problems, then the effectiveness of such treatments in preventing the progression to depression should also be assessed. From an aetiological perspective, understanding the mechanisms that lead children with conduct problems to become depressed would become a research priority. To examine associations between conduct problems and depression we use data from the Avon Longitudinal Study of Parents and Children (ALSPAC), a large population-based cohort which has a number of advantages for a study of this kind, including information about conduct problems starting early in childhood (4 years of age), data on potential risk factors, and multiple assessments from pregnancy through to young adulthood. We sought to answer five main questions:

(a) What is the risk for adult depression in children with early-life conduct problems, and how does this risk compare with that of early-life depression? Moreover, is the association between conduct problems and depression robust to adjustment for early depression?

(b) Do some subtypes of conduct problems carry a particular risk for early adult depression? We focused on age at onset as a key differentiating factor, in light of extensive research into the hypothesis that disruptive behaviours emerging in childhood differ from adolescent-onset conduct problems in both aetiology and course.8,9

(c) Do young people with a history of conduct problems show the typical adolescent rise10 in depressive symptoms?

(d) Are depressive symptom profiles similar in young people with depression with and without a history of conduct problems? We anticipated that irritability would be more common among young adults with depression and a history of conduct problems, based on the finding11 that conduct problems are more common in adolescents with depression who are also irritable.

(e) Although much past attention has focused on adolescent risk factors for depression in teenagers with conduct problems,12 little research has focused on the role of early childhood factors. The rich early-life data available in the ALSPAC cohort allow us to examine the role of a range of potential confounders.

Sample
ALSPAC is an ongoing population-based study designed to investigate the effects of a number of factors on health and development. All women resident in Avon, UK, with expected dates of delivery between 1 April 1991 and 31 December 1992 were contacted and eligible for participation. The study cohort consisted of 14,541 pregnancies and 13,988 children still alive at 12 months of age. Compared with the 1991 UK national census data (available at http://www.ons.gov.uk), the sample showed a slightly higher proportion of homeowners and a smaller proportion of mothers from Black and minority ethnic groups. As described by Boyd et al.13 children enrolled in ALSPAC were more educated at 16 compared with the national average, were...
more likely to be White (reflecting the ethnic composition of the area from which the sample was drawn) and less likely to be eligible for free school meals (an indicator of low income). Ethical approval for the study was obtained from the ALSPAC Law and Ethics Committee and local research ethics committees. Detailed information has been collected on the cohort since early pregnancy, including regular self-reported information from the mothers, and child assessments made in research clinics. The study website contains details of all the data that are available through a fully searchable data dictionary: www.bris.ac.uk/alspac/researchers/data-access/data-dictionary.

Measures
Young adult depression
Young adult depression was assessed using the Revised Clinical Interview Schedule (CIS-R),14 a self-administered, computerised interview completed at age 18 at the research clinic (mean age at attendance 17 years 10 months). The CIS-R establishes the severity of core symptoms of depressive disorders (depression, depressive thoughts, fatigue, sleep and concentration problems), along with diagnostic cut-points for depressive disorders, using algorithms based on ICD-10 criteria.15 Each symptom is scored on a 0–4 scale (depressive thoughts 0–5), according to the severity (frequency, duration and unpleasantness) of the symptom experienced. The CIS-R also provides indicators of the severity of irritability (scored 0–4), anhedonia (scored 1–3) and suicide risk (scored 0–3), and the frequency of self-blame (scored 1–4). A depression severity score based on the sum of the ICD-10 symptom criteria in the CIS-R (depression, depressive thoughts, sleep and concentration problems and fatigue; range 0–21, Cronbach’s \( \alpha = 0.77 \)) was used as an adjustment in analyses of individual depression symptoms (see below).

Depressive symptoms in late childhood and adolescence
Depressive symptoms in late childhood and adolescence were assessed using the short (13-item) Mood and Feelings Questionnaire (MFQ)16,17 to enquire about the occurrence of depressive symptoms over the past 2 weeks. The MFQ was completed by ALSPAC study members at research clinics at ages 10.5 and 13 years and by postal questionnaire at 16 years; slight variations in actual ages at questionnaire completion were unrelated to levels of depressive symptoms. The internal construct validity of the MFQ has been extensively supported, including in a UK community sample, which confirmed a single continuum of severity of depressive symptoms.18 The internal reliabilities (Cronbach’s \( \alpha \)) of the items in the ALSPAC sample were 0.80, 0.86 and 0.91 at ages 10, 13 and 16 respectively. As expected, MFQ scores varied by gender across the three time points, with higher scores among boys at age 10, and among girls at ages 13 and 16 (age 10, boys: mean = 4.05 (s.d. = 3.38), girls: mean = 3.85 (s.d. = 3.51), \( b = -0.19, 95\% \) CI \(-0.37 \) to \(-0.02), \( P < 0.05 \); age 13, boys: mean = 4.07 (s.d. = 3.80), girls: mean = 5.60 (s.d. = 4.79), \( b = 1.53, 95\% \) CI \(1.29 \) to \(1.78, P < 0.001\); age 16, boys: mean = 4.26 (s.d. = 4.57), girls: mean = 6.73 (s.d. = 5.82), \( b = 2.47, 95\% \) CI \(2.15 \) to \(2.78, P < 0.001\)). For the purpose of estimating population attributable fractions (PAFs; see below), total MFQ scores at ages 10 and 13 were dichotomised at \( \geq 90\% \) percentile as a ‘high’ cut-off.

Conduct problems
Maternal reports of child conduct problems were collected at ages 4, 7, 8, 10, 12 and 13 years using the Strengths and Difficulties Questionnaire (SDQ), a widely-used screening instrument with well-established reliability and validity.19,20 The conduct problem subscale of the SDQ includes five items designed to assess a broad spectrum of disruptive behaviours including fighting, lying, stealing, disobedience and temper outbursts. Four developmental trajectories of conduct problems were identified in previous analyses using growth mixture models;21 an early-onset persistent class (EOP, 7.8% of the present analytic sample, 56.4% boys), who initiated conduct problems early in childhood and persisted into the early teens; a childhood-limited class (CL, 14.0% of sample, 52.9% boys), who initiated conduct problems in early childhood but desisted to low levels by age 13; an adolescent-onset class (AO, 10.7% of sample, 51.1% boys), who started at low levels of conduct problems in childhood but increased by age 13; and a ‘low’ class (low, 67.5% of sample, 48.8% boys), who maintained low levels of conduct problems throughout. For some analyses, we combined the three ‘high’ conduct problem trajectory classes (EOP, CL and AO) to create a ‘conduct problems ever’ category, and contrasted this with the ‘low’ class.

Early family, parental and sociodemographic adversity
Multiple indicators of family, parental and sociodemographic risk were assessed during pregnancy via a Family Adversity Index (FAI).22 The FAI consisted of 17 items (including maternal age, education and marital status, housing and financial difficulties, partner relationships and support, parental psychopathology, substance use and crime) taken from questionnaires administered throughout the mother’s pregnancy (8, 12, 18 and 32 weeks gestation). If adversity was present, an item was rated as 1 and scores were summed to create a scale (range 0–10). Just over half of cases (52.6%) in the present sample were not exposed to any of these adversities, 28.2% to one, and 19.3% to two or more.

Child characteristics
Temperament was assessed using the emotional subscale of the Emotionality, Activity, Sociability (EAS) Temperament questionnaire (score range 0–20, mean = 7.71, s.d. = 3.34),23 completed by mothers when the study children were 69 months old. Reading was assessed using the WORD24 test of reading accuracy (score range 0–52, mean = 28.78, s.d. = 9.05) administered by trained psychologists in research clinics at age 7.5 years. Self-esteem was assessed at age 8 at the research clinics using the scholastic competence subscale of Harter’s Self-Perception Profile for Children25 (score range 6–24, mean = 17.09, s.d. = 3.62).

Attrition
Data to estimate conduct problem trajectories were available for 7218 young people. As detailed in past reports,21 cases with missing data for trajectory assignments differed from the full birth cohort on a number of sociodemographic background factors (including higher rates of teen motherhood and maternal smoking in pregnancy, and lower levels of maternal education). Age 18 depression data were available for 3542 individuals with conduct problem trajectory assignments. Girls were more likely than boys to participate at age 18 (54.7% v. 43.5%; Pearson \( \chi^2 (1) = 90.25, P < 0.001\)). Adjusted for gender, trajectory classes were also significant predictors of participation at age 18, with odds ratios (ORs) ranging from 0.83 (95% CI 0.76–0.91, \( P < 0.001\)) for those in the CL class to 0.69 (95% CI 0.60–0.80, \( P < 0.001\)) for those in the EOP class. To ensure the robustness of our findings we repeated the study’s main analyses using inverse probability weighting (see online supplement for details) and found essentially the same pattern of results; we therefore present the unweighted, full case analyses in this paper.
Statistical methods

Conduct problem trajectories and young adult depression

We estimated associations between age 18 depression and either conduct problems ever or conduct problem trajectory classes using logistic regression models, with age 18 depression diagnosis as outcome and conduct problems ever or trajectory classes as predictors. To compare the estimates obtained from the various conduct problem indicators with those for depression, the dichotomised MFQ scores were entered either on their own or together with conduct problems ever or conduct problem trajectory classes in logistic regression models. The PAF for the predictors was estimated in Stata version 11.2 for Windows using the punaf command. Unless otherwise indicated, the low conduct problem trajectory class was used as the reference category. We present odds ratios, 95% confidence intervals and P-values from these analyses.

Conduct problems and the development of depressive symptoms in adolescence

The relationship between conduct problem trajectories and MFQ scores at age 10 years was examined in a linear regression model, with trajectories as predictors; we present b-coefficients with 95% confidence intervals and P-values. To examine associations between trajectories and the development of depressive symptoms across adolescence, we estimated the baseline levels (‘intercept’ at age 10) and rate of change (‘slope’ between ages 10 and 16) of MFQ scores in Mplus.26 We imported these estimates into Stata and used them in linear regression models with the slope as the outcome and trajectory classes as the predictors, adjusting for baseline (age 10) level of depression; we present regression coefficients with 95% confidence intervals and P-values.

Depressive symptom profiles

Associations between trajectory classes and depression symptoms at age 18 were estimated in ordinal logistic regression models with symptom scores as outcomes, adjusted for the overall depression severity score based on the sum of the ICD-10 symptom criteria in the CIS-R. We present odds ratios, 95% confidence intervals and P-values from these analyses.

Early childhood confounds

These were entered in logistic regression models with age 18 depression diagnosis as outcome and conduct problems ever and conduct problem trajectory classes as predictors. We present odds ratios, 95% confidence intervals and P-values from these analyses.

Unless otherwise specified, all analyses were undertaken in Stata version 11.2 for Windows, with cases weighted by the posterior probabilities of trajectory group membership throughout.

Results

Childhood conduct problems and young adult depression

There was a strong association between young adult depression and both conduct problems ever and the conduct problem trajectory classes, as illustrated in Fig. 1. Overall, 42.3% of those with depression at age 18 had shown high levels of conduct problems at some point in childhood or early adolescence. As expected, girls were more likely to meet depression criteria than boys (OR = 2.54, 95% CI 1.98–3.27, P < 0.001). Adjusted for these strong gender effects, both conduct problems ever and conduct problem trajectory classes remained significantly associated with risk for depression at age 18 (Table 1). There was no evidence of a gender x trajectory class interaction (F(3,6058) = 0.47, P = 0.70). In addition, young people with early-onset and persistent conduct problems were at significantly higher risk of depression at age 18 than those in the CL (OR = 0.58, 95% CI 0.40–0.82, P = 0.002) and AO (OR = 0.65, 95% CI 0.46–0.92, P = 0.014) classes.

We next compared the following in terms of their association with young adult depression: (a) conduct problems ever or conduct problem trajectory classes; (b) depression in late childhood or early adolescence (age 10 and 13 MFQ scores).

Table 1 presents the odds ratios for the association between each predictor and outcome adjusted for gender, and also the PAFs for the predictors of interest. The PAF estimate for adult depression due to conduct problems ever was about twice as large as that for depression at age 10 years, and comparable to that for depression at age 13 years (15% vs. 19%). Finally, we estimated the PAF due to conduct problems adjusting for high depression scores at ages 10 and 13 years. As can be seen in Table 1, the PAF estimates changed only slightly (from 0.15 to 0.14) following the adjustment.

Conduct problem trajectories and the development of depressive symptoms in adolescence

Figure 2 illustrates the development of depressive symptoms (as assessed by mean total MFQ scores) in each gender and trajectory class from late childhood to middle adolescence. At age 10 years there was a significant effect of trajectory class, with young people in each of the ‘conduct problem’ trajectories showing significantly higher symptom levels than those in the low class (CL: b = 0.61, 95% CI 0.42 to 0.81; AO: b = 0.84, 95% CI 0.64 to 1.05; EOP: b = 1.62, 95% CI 1.30 to 1.95; all P < 0.001). There was no significant gender x conduct problem trajectory class interaction (F(3,6709) = 2.29, P = 0.08). We next estimated the intercept (mean for boys = 3.84 s.d. = 1.22; mean for girls = 4.07, s.d. = 1.34) and slope (mean for boys 0.70, s.d. = 0.70; mean for girls 0.91, s.d. = 0.88) for depressive symptoms between ages 10, 13 and 16 years. As expected, there was a considerable increase in depressive symptom levels between ages 10 and 16 years (Fig. 2), reflected in the results from a linear growth model (estimate for slope = 0.90, 95% CI 0.82 to 0.98, P < 0.001). In a model estimating the rate of change (slope) in depressive symptoms across ages 10, 13 and 16 years, gender was a significant predictor, with girls showing stronger increases in symptom scores than boys (b = 0.30, 95% CI 0.26 to 0.33, P < 0.001). As Fig. 2 suggests, however, there was no difference in the rate of change in depressive symptoms by conduct problem trajectory after gender and baseline (age 10) symptom levels.
had been taken into account (CL: b = 0.01, 95% CI = 0.02 to 0.05; AO: b = 0.03, 95% CI = 0.01 to 0.07; EOP: b = 0.02, 95% CI = 0.04 to 0.07; all P > 0.15). There was also no significant gender × conduct problem trajectory interaction in the prediction of the rate of symptom change (F(3,6939) = 0.51, P = 0.67).

### Depressive symptom profiles

Figure 3 focuses on young people who met criteria for depression at age 18, and shows mean symptom severity scores for each conduct problem trajectory class. Preliminary tests showed no evidence of gender differences in symptom severity within individuals meeting criteria for depression (b = 0.59, 95% CI = 0.34 to 1.51, P = 0.214), so data for males and females are combined. In general, as Fig. 3 suggests, depressive symptom profiles were very similar irrespective of conduct problem history. In keeping with our hypothesis, however, irritability was significantly more common in the EOP (OR = 2.40, 95% CI 1.34 to 4.31, P = 0.004) and CL (OR = 1.59, 95% CI 1.03 to 2.45, P = 0.04) classes than in the low class; young people classified as showing adolescent-onset conduct problems did not differ significantly from the low group in levels of irritability (OR = 1.29, 95% CI 0.77 to 2.16, P = 0.33). After adjusting for overall depression severity, irritability was still significantly more common in the EOP (OR = 2.16, 95% CI 1.24 to 3.78, P = 0.007) than in the low class; adjustment for depression severity

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**Table 1** Predictors of young adult depression by conduct problems ever, conduct problem trajectories and early depression

<table>
<thead>
<tr>
<th>Conduct problems ever</th>
<th>Conduct problem trajectory class</th>
<th>Late childhood/early adolescent depression (&gt; 90th percentile MFQ score)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>CL</td>
<td>AO</td>
</tr>
<tr>
<td>Adjusted for gender</td>
<td></td>
<td></td>
</tr>
<tr>
<td>OR</td>
<td>1.62 (1.30–2.01)</td>
<td>1.34 (1.03–1.73)</td>
</tr>
<tr>
<td>PAF</td>
<td>0.15 (0.08–0.22)</td>
<td>0.04 (0.00–0.07)</td>
</tr>
<tr>
<td>PAF</td>
<td>0.07 (0.01–0.11)</td>
<td>0.00 (0.00–0.07)</td>
</tr>
</tbody>
</table>

MFQ, Mood and Feelings Questionnaire; CL, childhood limited; AO, adolescent onset; EOP, early-onset persistent; OR, odds ratio; PAF, population attributable fraction.

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**Fig. 2** Mean Mood and Feelings Questionnaire (MFQ) scores at ages 10, 13 and 16 years by conduct problem trajectories and gender.

**Fig. 3** Mean age 18 depression symptom scores by conduct problem trajectories: individuals meeting criteria for depression only.
Conduct problems ever
Adjusted OR\(^a\) (95% CI) Conduct problem trajectory class
Baseline model (reference category for trajectory class comparisons = ‘low’)
Baseline model
1.62 (1.30–2.01) CL Adjusted OR\(^a\) (95% CI)
Adjusted for:
Family Adversity Index 1.48 (1.19–1.83) 1.34 (1.03–1.73) 1.50 (1.16–1.94) 2.32 (1.64–3.27)
Child characteristics
Temperament (emotionality, age 5 years) 1.55 (1.23–1.95) 1.27 (0.98–1.64) 1.40 (1.08–1.81) 1.99 (1.40–2.62)
Reading (age 7 years) 1.69 (1.36–2.11) 1.29 (0.99–1.68) 1.42 (1.08–1.86) 2.21 (1.55–3.14)
Academic self-esteem (age 8 years) 1.52 (1.23–1.90) 1.44 (1.10–1.87) 1.52 (1.17–1.99) 2.51 (1.75–3.62)
All significant confounders\(^b\)
1.50 (1.18–1.89) 1.30 (0.98–1.71) 1.40 (1.06–1.86) 2.08 (1.40–3.08)

Table 2 Associations between conduct problems and early adult depression: effects of adjustment for early childhood confounds

\(a\) All odds ratios (ORs) adjusted for gender.

\(b\) Family Adversity Index, age 7 reading, age 8 academic self-esteem.

Discussion

In this paper we set out to address five key outstanding questions concerning the extent and nature of the overlap between conduct problems in childhood and early adult depression.

The first was to address current ambiguities in the relationship between conduct problems in youth as a precursor of adult depression, with some studies suggesting an association,\(^3\)–\(^5\) but others failing to find a link.\(^6\) It is probably as a result of this uncertainty that, despite increasing interest in early antisocial behaviours, depression is only rarely addressed as one of its outcomes.\(^7\) In our study using a large birth cohort, we confirm that conduct problems were strong markers of risk for subsequent depression: over 43% of young adults with depression in the ALSPAC cohort had a history of conduct problems in childhood or adolescence. Our study also demonstrates the potential public health importance of these findings by showing that PAFs for adult depression due to conduct problems were comparable to those due to early adolescent depression. This would imply that a considerable number of cases of young adults with depression could be prevented if early conduct problems were treated adequately and the association were causal. Effective treatments for conduct problems in youth\(^7\) might prevent some future depression, a hypothesis that our findings suggest should be tested empirically. Moreover, our study shows that the prediction from early conduct problems to young adult depression is largely independent of depression at ages 10 and 13 years. Of note, the prediction from conduct problems to young adult depression is cross-informant (from early parent- to later self-report), adding to the robustness of our findings.

Against this background, we turned to our second question: whether particular subtypes of conduct problems might carry an especially high depression risk. We focused on variations in age at onset, in light of extensive research on their roles as key markers of heterogeneity in antisocial behaviour.\(^8\) More recently, distinctions have also been drawn within early-onset cases between individuals that persist in antisocial activities and those that do not.\(^9\) To date, however, few studies have examined variations among these subtypes in risk for early adult depression. One possibility is that early-onset conduct problems (which typically show poor later outcomes) would also carry the strongest risk for depression; another is that conduct problems may be more strongly associated with depression if they are present in adolescence, when depressive symptoms themselves show a sharp increase.\(^27\)

Of the few tests of these possibilities reported thus far, Odgers et al.,\(^9\) using data from the Dunedin longitudinal cohort, provide the closest match to our characterisation of child and adolescent conduct problems. They examined past-year rates of major depressive disorder in adulthood (age 32) for individuals following child and adolescent antisocial trajectories broadly similar to those examined here. In both men and women, risks of depression in the early 30s were markedly elevated for those on a ‘life course persistent’ antisocial trajectory, but were not significantly elevated for those on either childhood- or adolescence-limited pathways. Our findings were consistent with this pattern in finding much elevated rates of depression among young people with early-onset and persistent conduct problems, significantly higher than those in all other conduct problem trajectory groups. In addition, however, our findings suggested elevated risks for young people with both childhood-limited and adolescent-onset conduct problems. It is unclear at this stage how far these variations reflect differences in the ages at which
family adversities, child temperament, early cognitive skills and
self-esteem resulted in only modest reductions in estimates of risk
for young adult depression associated with childhood conduct
problems. Our study cannot determine whether these associations
reflect a direct causal impact of conduct problems on risk for
depression or the operation of other shared risk factors, common
to both disorders. It has been argued, for example, that the adverse
correlations of conduct problems for children’s family and peer
relationships, and for their skill development and academic
performance result in a ‘failure trajectory’ that increases
vulnerability to depression.31 In addition, however, subsequent
research suggests that the relationship between depression and
conduct problems can largely be accounted for by shared genetic
factors.28,32 Further research on these differing pathways would
clearly be of importance.

Strengths and limitations
Our study has several strengths, including a large sample size, a
longitudinal design with rich measures spanning childhood to
young adulthood, and data from multiple informants. As is
common in longitudinal studies, however, there was loss to
follow-up over the period covered by our analyses, and it was clear
that these losses were selective: more socially disadvantaged
individuals were excluded from the trajectory group assignments,
and conduct problems were themselves significant predictors of
study member participation in early adulthood. As a result, we
are likely to have underestimated the prevalence of conduct
problems in our analytic sample. Past evidence suggests, however,
that it is much less likely that loss to follow-up would affect
estimates of the associations between variables,31 and the use of
inverse probability weights produced essentially similar results in
the present analyses; as a result, we think it unlikely that the
associations we observed between conduct problem trajectories
depression are explained by sample attrition.

Clinical implications
Our findings have a number of implications. First, in keeping with
past findings, the data from this study suggest that conduct
problem trajectories are a marker of risk for depression problems
throughout adolescence and early adulthood; hence, clinicians
treating conduct problems in children and adolescents should be
alert to recognise symptoms of depression and not to miss early
treatment opportunities. Second, our findings suggest that a
history of childhood or adolescent conduct problems, and not
solely the persistence of such problems, carries risk for the
development of depression. From a prevention and treatment
perspective, it is important to note that even those in the CL group
– i.e. those children whose conduct problems had desisted by
development – were more likely than the rest of the population
to have depression in early adulthood. These findings may suggest that
interventions aimed at depression, over and above those aimed at
targeting disruptive behaviour problems, may be necessary in
children presenting with conduct problems. Future studies
investigating the mechanisms underlying these transitions, and
how these could inform early interventions, would be of value.

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We are extremely grateful to all the families who took part in this study, the midwives for their help in recruiting them, and the whole ALSPAC team, which includes interviewers, computer and laboratory technicians, clerical workers, research scientists, volunteers, managers, receptionists and nurses. The UK Medical Research Council and the Wellcome Trust (Grant ref. 092733) and the University of Bristol provide core support for ALSPAC.

Incorporating inverse probability weights

To take some account of sample attrition we derived inverse probability weights\(^{35}\) from a logistic regression model predicting having complete conduct problems trajectory and age 18 depression data, adjusting for exposure to early life adversity (Family Adversity Index score) and child gender. The weights from these analyses were then multiplied by the posterior probability weights for trajectory class membership, with higher values trimmed to a maximum weight of 4. Table DS1 shows results of logistic regression analyses of associations between conduct problems trajectory classes (reference category = ‘low’) and age 18 depression (i) using posterior probability weights only, and (ii) using posterior probability weights × inverse probability weights. Child gender (not shown) was included as a covariate throughout.

As Table DS1 shows, models incorporating inverse probability weights provided essentially similar results to those using posterior probability weights only.

**Additional reference**


---

### Table DS1  Associations between conduct problem trajectories and age 18 depression: effects of incorporating inverse probability weights

<table>
<thead>
<tr>
<th>Weighted by</th>
<th>CL Adjusted OR(^a) (95% CI)</th>
<th>AO Adjusted OR(^a) (95% CI)</th>
<th>EOP Adjusted OR(^a) (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Posterior probability weights</td>
<td>1.34 (1.03–1.73)</td>
<td>1.50 (1.16–1.94)</td>
<td>2.32 (1.64–3.27)</td>
</tr>
<tr>
<td>Posterior probability × inverse probability weights</td>
<td>1.39 (1.05–1.86)</td>
<td>1.60 (1.20–2.13)</td>
<td>2.46 (1.75–3.46)</td>
</tr>
</tbody>
</table>

CL, childhood limited; AO, adolescent onset; EOP, early onset persistent.

\(^a\) All odds ratios adjusted for gender.
Developmental pathways from childhood conduct problems to early adult depression: findings from the ALSPAC cohort

Argyris Stringaris, Glyn Lewis and Barbara Maughan

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