Clinical and social factors associated with attention-deficit hyperactivity disorder medication use: population-based longitudinal study

Cédric Galéria,* Jean-Baptiste Pingault,* Grégory Michel, Manuel-Pierre Bouvard, Maria Melchior, Bruno Falissard, Michel Boivin,** Richard E. Tremblay** and Sylvana M. Côté**

Background
The impact of longitudinal psychiatric comorbidity, parenting and social characteristics on attention-deficit hyperactivity disorder (ADHD) medication use is still poorly understood.

Aims
To assess the baseline and longitudinal influences of behavioural and environmental factors on ADHD medication use.

Method
Survival regressions with time-dependent covariates were used to model data from a population-based longitudinal birth cohort. The sample (n = 1920) was assessed from age 5 months to 10 years. Measures of children's psychiatric symptoms, parenting practices and social characteristics available at baseline and during follow-up were used to identify individual and family-level features associated with subsequent use of ADHD medication.

Results
Use of ADHD medication ranged from 0.2 to 8.6% between ages 3.5 to 10 years. Hyperactivity–inattention was the strongest predictor of medication use (hazard ratio (HR) = 2.75, 95% CI 2.35–3.22). Among all social variables examined, low maternal education increased the likelihood of medication use (HR = 2.09, 95% CI 1.38–3.18) whereas immigrant status lowered this likelihood (HR = 0.40, 95% CI 0.17–0.92).

Conclusions
Beyond ADHD symptoms, the likelihood of receiving ADHD medication is predicted by social variables and not by psychiatric comorbidity or by parenting. This emphasises the need to improve global interventions by offering the same therapeutic opportunities (including medication) as those received by the rest of the population to some subgroups (i.e. immigrants) and by diminishing possible unnecessary prescriptions.

Declaration of interest
M.-P.B. has received financial support for the organisation of scientific meetings and was also the principal investigator in clinical trials for Shire and Lilly. In the past, C.G. received support from the industry to attend scientific congresses.

The first-line pharmacological intervention in attention-deficit hyperactivity disorder (ADHD) is stimulant medication, particularly methylphenidate. It has shown short-term efficacy in reducing behavioural symptoms of ADHD1,2 and improving cognitive tasks and academic performance.3,4 However, although use of medication for ADHD is associated with a possible decrease in criminality,5 there is little evidence of its impact on long-term impairments related to risky behaviours, psychiatric comorbidities and socio-occupational outcomes.6,7 Moreover, ADHD medication is generally well tolerated but potentially exposes children to adverse effects on appetite, growth, sleep and the cardiovascular system.8,9 In some cases ADHD medication is not adequately used since children without ADHD (false positives) or children with moderate forms can unduly be exposed to medication. Conversely, many children with severely impairing ADHD do not receive medication, even in countries with the highest rates of prescription.10 Accordingly, both expert assessment of diagnosis11,12 and expert monitoring and management of prescription practices are required.8 The international situation regarding ADHD medication prescription is variable.13 In many high-income countries, pharmacoepidemiological surveys suggest an increase in the pattern of stimulant medication use over time.14 This is notable in the USA, where the overall prevalence of medicated children rose from 0.6% in 1987 to 2.7% in 1997 and to 3.5% in 2008, although the situation varies from state to state (over- and underuse).15,16 However, in other countries, there is insufficient access to pharmacotherapy.

The substantial impact of ADHD on health and quality of life calls for a better understanding of its healthcare determinants, particularly medication. This is pressing, as alternative non-pharmacological interventions for ADHD (dietary and psychological treatments) have shown limited benefit for reducing core ADHD symptoms.17 Accordingly, medication still appears to be one of the most effective approaches to treating ADHD. However, there are risks of over-, under- and inadequate prescriptions. In addition, the concerns about the rising number of prescriptions, misuse and long-term safety have triggered worries among the stakeholders (the families of treated children, the media and the medical community), which may influence practices and acceptance of interventions.18,19 Identifying factors that predict medication use beyond ADHD symptoms is of utmost importance as this knowledge could be used by clinicians in their treatment decisions.

Prior research found that a variety of individual characteristics, sociodemographic and environmental factors are associated with the prescription of medication for ADHD: youth psychopathology (ADHD, externalising disorders, internalising disorders), being male, ethnicity, non-intact families, parental psychopathology, low maternal education, family income, negative family influences, low academic functioning, and previous receipt of stimulant medication.16,20–26 For some predictors, particularly disruptive comorbidity, divergent findings have been reported (protective

*These authors contributed equally to the work; **these authors share the senior authorship.
in some studies v. risk factor in others). In addition, these studies had limitations: (a) medication use was often measured only at one point in time; (b) some studies were cross-sectional, implying a concomitant evaluation of risk factors and outcomes; (c) potential confounders were not taken into account; (d) a categorical diagnosis of psychiatric problems was generally used; (e) early childhood factors were not studied prospectively in a population-based sample. These limitations impeded the correct longitudinal appraisal of the role of risk factors, exposed the data to confounding biases and did not allow a dimensional approach to behavioural problems. The aim of the present study was to go beyond these limits by using a population-based birth cohort to assess the baseline and longitudinal influences of environmental and behavioural predictors on ADHD medication between the ages of 3.5 and 10 years. We tested the hypotheses that: (a) hyperactivity–impulsivity symptoms and inattention symptoms would predict medication use in this population-based sample; (b) other risk factors could heighten/lessen the likelihood of receiving ADHD medication.

**Method**

**Participants and procedure**

Data were drawn from the Quebec Longitudinal Study of Child Development (QLSCD). The QLSCD protocol was approved by the Quebec Institute of Statistics and the St Justine Hospital Research Center ethics committees. Data were collected by trained interviewers through repeated (n = 10) home interviews with the person most knowledgeable about the child (the mother for 98% of children) in order to obtain information on child, parent and family characteristics and behaviours. Informed written consent was obtained from all participating families at each assessment. Assessments were conducted at the following ages: 5 months, 1.5, 2.5, 3.5, 4, 5, 6, 7, 8 and 10 years. The initial sample was selected from birth registries and comprised 2120 children evaluated at 5 months and representative of children born in the province of Quebec (Canada) in 1997/1998. The average response rate over the 10 years of data collection was 83% (range: 63–100%) with an average completeness of data equal to 79% (range: 61–91%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%). The sample with complete data at the first response rate over the 10 years of data collection was 83% (range: 63–100%).

**Measures**

**Outcome:** use of ADHD medication

Use of ADHD medication was reported by parents in a question referring to the preceding 12 months ‘Does [your child] take any of the following prescribed medication on a regular basis: Ritalin or any other medication for treating hyperactivity or inattention?’ at ages (years) 3.5 (0.2% of the sample), 4 (0.2% of the sample), 5 (0.2% of the sample), 6 (1.5% of the sample), 8 (5.6% of the sample), and 10 (8.6% of the sample).

**Explanatory variables**

**Time-varying covariates: children’s mental health and parenting.** Symptoms of hyperactivity–impulsivity, inattention, anxiety, opposition and emotional problems were reported through the Interviewer Computerized Questionnaire when the children were 2.5, 3.5, 4, 5, 6, and 8 years of age. Ratings relied on the early childhood behaviour scale from the Canadian National Longitudinal Study of Children and Youth. This tool incorporates items from the Child Behavior Checklist, the Ontario Child Health Study Scales and the Preschool Behavior Questionnaire. Each dimension the items used were as follows:

(a) hyperactivity–impulsivity: ‘could not sit still, was restless or hyperactive’, ‘could not stop fidgeting’, ‘was impulsive, acted without thinking’, ‘had difficulty waiting for his/her turn in games’, ‘couldn’t settle down to do anything for more than a few moments’;

(b) inattention: ‘was unable to concentrate, could not pay attention for long’, ‘was easily distracted, had trouble sticking to any activity’, ‘was inattentive’;

(c) anxiety: ‘was too fearful or anxious’, ‘was worried’, ‘was nervous, high strung or tense’, ‘cried a lot’;

(d) opposition: ‘was defiant or refused to comply with adults’ requests or rules’, ‘didn’t seem to feel guilty after misbehaving’, ‘punishment didn’t change his/her behaviour’, ‘had temper tantrums or hot temper’;

(e) emotional problems: ‘seemed to be unhappy or sad’, ‘was not as happy as other children’, ‘had no energy, was feeling tired’, ‘had trouble enjoying him/herself’.

All items referred to the preceding 12 months and were coded on a frequency scale (never or not true: 0; sometimes or somewhat true: 1; often or very true: 2) and quantitative scores derived from scales were z-standardised. Owing to high correlations and to avoid multicollinearity, hyperactivity–impulsivity and inattention were combined into a single variable hyperactivity–inattention.

The Parental Cognition and Conduct Toward the Infant scale was used to assess coercive parenting when the children were 2.5, 3.5, 4, 5, 6, and 8 years of age, using the following items: ‘I have been angry with my child when he/she was particularly fussy’, ‘when my child cries, he/she gets on my nerves’, ‘I have raised my voice with, or shouted at, my child when he/she was particularly fussy’, ‘I have spanked my child when he/she was particularly fussy’, ‘I have lost my temper when my child was particularly fussy’, ‘I have left my child alone in his/her bedroom when he/she was particularly fussy’, ‘I have shaken my child when he/she was particularly fussy’. All items were rated on an 11-point scale (higher score, more coercive parenting).

**Baseline covariates**

The gender of the child was coded 1 for boys and 0 for girls. Family structure was coded 1 if the family was non-intact (i.e. single-parent families; or families composed of a couple, married or common-law, living with at least one child not born to them) and 0 if the family was intact (i.e. the child lives with his/her two biological parents regardless of the type of conjugal relationship). Maternal education was coded 1 if low (no high-school diploma) and 0 if medium (high-school or post-secondary diploma) or high (university degree). Maternal age at birth of the target child was coded 1 if 21 years or younger (10.5%) and 0 if older than 21 years. Household income was computed according to Statistics Canada’s guidelines accounting for the family zone of residence, the number of people in the household and the family income in the past year. Income was coded 1 if insufficient and 0 if sufficient. Parental immigration status corresponded to non-immigrant (mother and father born in Canada) v. immigrant (mother or father born outside Canada; 74% belonging to racial/ethnic minorities). Maternal and paternal depressive

Galéa et al
symptoms were assessed with the abbreviated version (12 items) of the Center for Epidemiological Studies Depression Scale (CES-D).32 Parents reported the frequency of depressive symptoms in the previous week. Each item was coded on a four-point scale. Informant total ratings were z-standardised. All baseline variables were assessed when the target child was 5 months of age.

**Statistical analyses**

Longitudinal risk factors (i.e., multiple measurements) were treated as time-varying covariates in all analyses. We first described the sample characteristics and the patterns of unadjusted associations between the main risk factors and ADHD medication use by computing the instantaneous hazard (equivalent to the proportion of the population using ADHD medication per unit time) on the basis of Kaplan–Meier survival functions. Second, we conducted bivariate and multivariate analyses using survival analysis with time-dependent covariates to determine associations between baseline/longitudinal risk factors and subsequent use of an ADHD medication. The method made it possible to estimate hazard ratios (HR) for the use of ADHD medication (endpoints at waves: 3.5, 4, 5, 6, 8, and 10 years of age) adjusting for the effects of other time-varying covariates (measured at waves: 2.5, 3.5, 4, 5, 6, and 8 years of age). To select predictors included in the multivariate models, we estimated bivariate associations between risk factors and the outcome (survival analysis). Variables with P < 0.05 were entered into the multivariate models. Each participant was specified as a cluster of correlated observations, as medication status at any given time can depend on previous medication status. A robust variance was thus estimated to account for this pattern. In order to test the robustness of the findings, sensitivity analyses were performed using multiple imputation models (number of imputation, 100) under the missing-at-random (MAR) non-response mechanism. Finally, interactions between independent variables kept in the final model were tested. A P < 0.05 was considered statistically significant.

**Results**

Table 1 summarises the characteristics of the sample at baseline and the number of person-years and events during follow-up. Figure 1 shows the instantaneous hazard rates of ADHD medication use: raw rates (Fig. 1(a)), by levels of hyperactivity–inattention symptoms (> +1 s.d., +1 s.d. to median, < median) (Fig. 1(b)), by levels of maternal education (low/medium/high) (Fig. 1(c)), and by parental immigration status (immigrant v. non-immigrant) (Fig.1(d)).

Table 2 provides the results of proportional hazards regressions with time-dependent covariates. It shows the hazard ratios and 95% confidence intervals for the use of ADHD medication. The fully-adjusted model (number of person-years observed, 9165; number of events, 210) was significant (Wald $\chi^2 = 88.1$, d.f. = 8, $P < 0.0001$). All variables respected the hazard proportional assumptions (the contribution of each predictor was constant over time) except family income, which was not significantly associated with the outcome in multivariate models.

To deal with this issue, we stratified on this variable, which still controls for the variable but does not provide an estimate of its contribution. Hyperactivity–inattention, male gender, low maternal education and immigration status were significantly associated with the use of ADHD medication. There was no statistically significant interaction. Additional analyses using multiple imputed data showed the same patterns of associations: hyperactivity–inattention ($P < 0.0001$), gender of the child ($P < 0.0001$), low maternal education ($P = 0.006$) and immigration status ($P = 0.036$) were significantly associated with the use of ADHD medication. Complementary analyses dichotomising hyperactivity–inattention into two dimensions hyperactivity–impulsivity and inattention showed that both dimensions were significantly associated with ADHD medication exposure (online Table DS1).

**Discussion**

Potential risks and possible misprescription have led some clinicians to highlight the need for caution in medicating ADHD.18 Our findings suggest that ADHD medication is significantly predicted by core ADHD symptoms but also over and above these symptoms by additional risk factors: being male, low maternal education and immigrant status.

**Comparison with previous findings and interpretation**

As expected, the strongest clinical predictors of ADHD medication use were symptoms of hyperactivity–inattention. This finding is reassuring since hyperactivity–inattention symptoms are the core behavioural symptoms of ADHD explicitly targeted by this medication. Interestingly, both dimensions (i.e. hyperactivity–impulsivity and inattention) were independently related to medication. Comorbid psychiatric symptoms assessed longitudinally (i.e. anxiety, emotional problems and opposition) were not independently associated with ADHD medication use when other risk factors were accounted for. In other words, these other symptoms (particularly symptoms of opposition that are often correlated with ADHD symptoms) did not induce medication. This means that clinicians did not medicate children who showed oppositional/anxious problems beyond what their ADHD status required. These results are particularly relevant in the context of public discussions regarding the ethics of medicating youths with psychotropic medications.
Fig. 1 Instantaneous hazards of attention-deficit hyperactivity disorder (ADHD) medication use.
(a) ADHD medication use; contribution of (b) hyperactivity–inattention; (c) maternal education; and (d) parental immigration status.

Table 2 Survival models predicting exposure to attention-deficit hyperactivity disorder (ADHD) medication*

<table>
<thead>
<tr>
<th></th>
<th>Unadjusted hazard ratio (95% CI)</th>
<th>Fully adjusted hazard ratio (95% CI)</th>
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<tbody>
<tr>
<td>Male gender</td>
<td>3.19 (2.13–4.79)*</td>
<td>2.14 (1.44–3.18)*</td>
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<tr>
<td>Psychiatric symptoms</td>
<td></td>
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<td>Hyperactivity–inattention</td>
<td>2.85 (2.50–3.26)*</td>
<td>2.75 (2.35–3.22)*</td>
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<tr>
<td>Anxiety</td>
<td>1.29 (1.09–1.52)*</td>
<td>0.96 (0.82–1.12)</td>
</tr>
<tr>
<td>Opposition</td>
<td>1.59 (1.35–1.86)*</td>
<td>0.85 (0.71–1.01)</td>
</tr>
<tr>
<td>Emotion</td>
<td>1.09 (0.94–1.27)</td>
<td>–</td>
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<tr>
<td>Coercive parenting</td>
<td>1.64 (1.36–1.97)*</td>
<td>1.19 (0.97–1.45)</td>
</tr>
<tr>
<td>Immigrant</td>
<td>0.28 (0.12–0.66)*</td>
<td>0.40 (0.17–0.92)*</td>
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<tr>
<td>Low maternal education</td>
<td>2.47 (1.65–3.69)*</td>
<td>2.09 (1.38–3.18)*</td>
</tr>
<tr>
<td>Non-intact family</td>
<td>1.89 (1.24–2.88)*</td>
<td>1.36 (0.86–2.15)</td>
</tr>
<tr>
<td>Insufficient family income</td>
<td>1.60 (1.07–2.40)*</td>
<td>–</td>
</tr>
<tr>
<td>Young maternal age at birth</td>
<td>1.56 (0.95–2.57)</td>
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<tr>
<td>Paternal depression</td>
<td>1.08 (0.92–1.28)</td>
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<td>Maternal depression</td>
<td>1.08 (0.90–1.31)</td>
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Data are courtesy of the Quebec Institute of Statistics.
*P < 0.05.
Boys were more likely than girls to receive medication, even when controlling for the frequency of hyperactivity–inattention symptoms. This association could be related to a diagnostic bias towards boys. This could be attributable to a noisier clinical expression of the disorder in boys and/or to cultural trends (mothering ideology and masculinity stereotypes). Consequently, boys could be more likely to benefit from a more precise ADHD diagnosis and treatment than girls.

Recent research on ADHD suggests that the disorder is a complex and aetiologically heterogeneous condition caused by a combination of genetic, environmental and epigenetic contributions. It has been argued that social and cultural characteristics may influence both diagnosis and prescription. The current study found associations consistent with this view. Most importantly and consistent with prior research, low maternal education was associated with ADHD prescription. There are several possible explanations for this finding. First, clinicians may conclude that psychoeducational interventions alone are less efficient than medication because of parental difficulties and limited resources. This would lead them to opt more easily for medication that has a quick and direct effect on ADHD symptoms and may make the parents more available for psychoeducative interventions. Second, mothers with more education may more readily seek and access information on medication, which makes them sensitive to drug issues and consequently less prone to accept use of medication. Third, low maternal education could confer risk for more severe symptomatology through a combination of environmental and biological factors. Fourth, phenomena such as parent-blame (more particularly mother-blame linked to the mothering ideology) and parental feelings of inadequacy in response to children's disruptive behaviour are thought to possibly bias ADHD diagnosis/treatment. All these potential social mechanisms need to be tested further by using qualitative and quantitative methods (for example randomised controlled trials) to examine the reasons for clinicians to suggest, and/or parents to accept, medication in families with different education levels, and examine the short- and long-term effectiveness of the different choices (for example to confirm or refute the benefit of more frequent prescriptions in lower-educated families). It is useful to note that other social variables were related to ADHD medication in univariate analyses (coercive parenting, non-intact family, insufficient family income) but were not associated with the study outcome in multivariate analyses, suggesting more distal influences or a less important role in medication use.

Parental immigrant status (which was strongly associated with racial/ethnic minorities) was related to lower ADHD medication use. This result is coherent with findings showing lower rates of ADHD diagnosis and less use of ADHD medication in children with a family immigrant background. The negative association could have several causes. First, immigrant status may be a barrier to proper access to healthcare. It has been suggested that minorities could be underdiagnosed owing to lesser healthcare utilisation because of poorer resources. Second, parents from ethnic minority groups have been shown to report fewer ADHD utilisation because of poorer resources. Clinicians may want to take parenting styles into account before prescribing and also may encourage clinicians to consider alternatives to drug prescription for this target population. Clinicians may want to take parenting styles into account before prescribing and also develop specific psychoeducational interventions dedicated to this population. Furthermore, the present study suggests that some subgroups with less access to their healthcare system and

**Strengths and limitations**

This study has several strengths, including the nature of the sample, the time period encompassed and the analytic approach. First, the large community-based sample made it possible to extend inferences beyond clinical populations. Second, the repeated measurement of ADHD and other behavioural symptoms prior to medication exposure increases reliability and strengthens causal inferences in comparison with prior cross-sectional studies. Third, since methylphenidate initiation most commonly begins between ages 5 to 9 years, this study provides an opportunity to study the predictors of prescription during this crucial period.

The study also has limitations. First, the study relies on parental reports, which makes the data subject to informant bias. Such reports might partly reflect higher or lower parental tolerance to ADHD behaviours. Of note, we adjusted for factors potentially associated with tolerance such as oppositional behaviours or non-intact family. Second, a full-blown categorical ADHD diagnosis was not a requirement. However, this limitation was offset by the use of a dimensional approach, thus providing a better assessment of the phenotype heterogeneity that leads to medication use. Third, there were no data on functional impairment related to ADHD symptoms. Future investigations should take impairment severity into account, in order to explore factors associated with prescriptions in moderate and severe forms of ADHD, which could help in understanding the issue of potential overdiagnosis and overtreatment. Fourth, information relating to non-pharmacological ADHD interventions was not available. Fifth, we could not consider potential mechanisms explaining the observed associations, for example cultural influences such as school pressure, parental expectations, community representations, mothering ideology, masculinity stereotypes and healthcare system representations. These should be examined in future work with randomised controlled trials in order to enhance our understanding of medication determinants. Finally, it is to be acknowledged that the influences of social variables on prescription depend on the environmental, cultural and legal context. The findings cannot be directly generalised to other country settings, especially newly industrialised nations and countries with lower prescription rates. Further research could benefit from cross-cultural designs and comparisons between different countries.

**Implications**

The study's findings have theoretical and practical implications. Physicians should bear in mind the possible contribution of social characteristics on their practices. Individual features of children are not the only variables influencing drug exposure. The role of social context (parental educational level, immigration status) and parenting style needs to be fully considered in order to choose effective therapeutic strategies. Caution needs to be taken since clinical decisions grounded solely on risk might stigmatise families and children on a long-term basis. However, within the set of possible explanations, both legitimate (i.e. accessibility to pharmacological and non-pharmacological interventions) and more questionable (i.e. social pressure) may link social variables to medication. An effort to design specific parenting and psychoeducative interventions for families with low educational levels may encourage clinicians to consider alternatives to drug prescription for this target population. Clinicians may want to take parenting styles into account before prescribing and also develop specific psychoeducational interventions dedicated to this population.
medication could benefit from selective interventions to enhance their access to psychosocial services and consequently be offered the same opportunities as the rest of the population (i.e. have equal access to good mental healthcare, defined as appropriate and unbiased use of ADHD medications).

Major practice guidelines 11, 12 recommend that the management of the ADHD clinical situation should rely on the joint use of psychoeducation, psychotherapy and pharmacotherapy. The current findings support the recommendations regarding the necessity of enhancing child and family therapeutic education and physicians’ knowledge of prescription determinants. Future research will need to provide a better understanding of parental and practitioners’ views to determine what drives decision-making on treatment choice.

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# Data supplement

<table>
<thead>
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<th>Table DS1</th>
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<td>1.89 (1.24–2.88)*</td>
</tr>
<tr>
<td>Insufficient family income</td>
<td>1.60 (1.07–2.40)*</td>
</tr>
<tr>
<td>Young maternal age at birth</td>
<td>1.56 (0.95–2.57)</td>
</tr>
<tr>
<td>Paternal depression</td>
<td>1.08 (0.92–1.28)</td>
</tr>
<tr>
<td>Maternal depression</td>
<td>1.08 (0.90–1.31)</td>
</tr>
</tbody>
</table>

Data are courtesy of the Quebec Institute of Statistics. *P < 0.05.
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Cédric Galéra, Jean-Baptiste Pingault, Grégory Michel, Manuel-Pierre Bouvard, Maria Melchior, Bruno Falissard, Michel Boivin, Richard E. Tremblay and Sylvana M. Côté
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