Mental health services, suicide and 7-day working
Nav Kapur, Saied Ibrahim, Isabelle M. Hunt, Pauline Turnbull, Jenny Shaw and Louis Appleby

Background
Patients admitted to hospital at the weekend appear to be at increased risk of death compared with those admitted at other times. However, a ‘weekend effect’ has rarely been explored in mental health and there may also be other times of year when patients are vulnerable.

Aims
To investigate the timing of suicide in high-risk mental health patients.

Method
We compared the incidence of suicide at the weekend v. during the week, and also in August (the month of junior doctor changeover) v. other months in in-patients, patients within 3 months of discharge and patients under the care of crisis resolution home treatment (CRHT) teams (2001–2013).

Results
The incidence of suicide was lower at the weekends for each group (incidence rate ratio (IRR) = 0.88 (95% CI 0.79–0.99) for in-patients, IRR=0.85 (95% CI 0.78-0.92) for post-discharge patients, IRR = 0.87 (95% CI 0.78–0.97) for CRHT patients). Patients who died by suicide were also less likely to have been admitted at weekends than during the week (IRR = 0.52 (95% CI 0.45–0.60)). The incidence of suicide in August was not significantly different from other months.

Conclusions
We found evidence of a weekend effect for suicide risk among high-risk mental health patients, but with a 12–15% lower incidence at weekends. Our study does not support the claim that safety is compromised at weekends, at least in mental health services.

Declaration of interest
N.K. is partly supported by Manchester Mental Health and Social Care Trust. N.K. chaired the guideline development group for the 2012 NICE guidelines on the longer-term management of self-harm and currently chairs the guideline development group for the NICE depression in adults guideline. L.A. chairs the Suicide Prevention Advisory Group at the Department of Health (of which N.K. is also a member) and is a non-executive Director for the Care Quality Commission.

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Data acquisition
Suicide data were collected as part of the National Confidential Inquiry into Suicide and Homicide by People with Mental Illness (NCISH). Data collection has been described in detail elsewhere. Briefly, it occurred in three stages. First, national data on people who had died by suicide in the general population in England were obtained from the Office of National Statistics. Second, NHS mental health services identified which of these individuals had been in contact with services in the 12 months before death. These trusts also identified the clinician who had been caring for the patient prior to their suicide. Third, detailed questionnaires were sent to the clinicians to obtain demographic and clinical data (including care at the time of death). Case ascertainment is relatively complete with a response rate for questionnaires of over 95%.

Deaths that received either a suicide or open verdict at coroner’s inquest were considered to be suicide deaths in the current study, as is the convention in suicide research in the UK. Not including open verdict deaths has been shown to underestimate the number of suicides by up to 50%. Both the cases that received a suicide verdict and those that received an open verdict are collectively referred to as ‘suicides’ in the rest of this paper.

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Study sample
The present study included individuals aged 10 years and over who died by suicide between 2001 and 2013 in England. The National Confidential Inquiry does not collect data on young children, where intentions and motivations may be difficult to establish.9 Previous studies of acute hospital mortality have used data collected over a single financial year,3 but because of the comparative rarity of suicide as an outcome we used data collected over a longer period. Some groups of patients with mental health disorders are at particularly high risk of suicide. They include psychiatric in-patients and those recently discharged from in-patient care,11,12 and those under the care of crisis resolution home treatment (CRHT) teams (a service intended as an alternative to admission).13 These groups are in close proximity to care and may be more affected by weekend changes to staffing or service availability than other patients. We therefore chose to focus on them in the current study.

Main outcome
Our main outcome was risk of suicide in relation to day of death. Previous investigations of the weekend effect in hospital settings have examined the risk of death by day of admission.5 Suicide commonly occurs in the context of complex difficulties, but the final act may be relatively impulsive and in response to an acute stressor.22 Care at the time of death may therefore be a more important determinant of outcome than care at the time of admission. In addition, lengths of stay for in-patient mental health services are much longer than for acute medical or surgical specialties,23 and so care immediately following admission may be less critical in terms of mortality risk.

Statistical analysis
We initially examined the timing of suicide by day of the week. We expressed the incidence of suicide as the number of suicide deaths per 100 days at risk (for example, the number of suicide deaths on a Monday per 100 Mondays at risk throughout the study period or the number of suicide deaths in January per 100 January days at risk). Suicide deaths are statistically rare events that can generally be expected to follow a Poisson distribution. Consequently, Poisson regression models were fitted with the number of suicides on each day as the dependent variable. Models were tested for overdispersion (where variation is high and violates the use of a Poisson model), and if this was evident, negative binomial regression models were fitted to account for high variation. The use of these models allowed the calculation of incidence rate ratios (IRR) with 95% confidence intervals, comparing the suicide incidence at the weekend with the suicide incidence during the working week. P-values less than 5% were considered significant. Levels of missing data were low—only two patients in the whole sample did not have details of the care they had been receiving at the time of death.

Although our main focus was on risk of suicide in relation to day of the week, we did also examine the risk of suicide by day of admission for the in-patient sample only. In order to investigate a possible August effect we examined the timing of suicide by month of the year. The ‘August effect’ (or its USA counterpart, the ‘July effect’) refers to the possible reduction in the quality and safety of care when final-year medical students become doctors and junior doctors become a grade more senior.17,24 There are other potential transition points, but the summer one is the best described and some of the other changeovers in the UK (for example every 4 months) are comparatively recent developments (post 2005).

Results
Over the study period there were 1621 in-patient suicide deaths, 2819 suicide deaths within 3 months of in-patient discharge and 1765 deaths under CRHT teams. Although in-patients were a distinct group, 592 (21%) of the post-discharge deaths were also under CRHT. Table 1 shows demographic and clinical characteristics of the patients.

Figure 1 shows the timing of suicide in relation to day of the week. There was a significant difference in the incidence of suicide between weekdays and weekends for all patient groups, with a lower suicide risk at weekends (in-patients: IRR = 0.88 (95% CI 0.79–0.99); post-discharge patients: IRR = 0.85 (95% CI 0.78–0.92); patients under CRHT: IRR = 0.87 (95% CI 0.78–0.97)).

Figure 2 shows the timing of suicide in relation to month of the year. There was no evidence of an August peak in suicide. The peak month for incidence of suicide was May for in-patients, September for post-discharge patients and November for CRHT team patients.

Table 1 Demographic and clinical characteristics of the patients who died by suicide between 2001 and 2013 in the study

<table>
<thead>
<tr>
<th>Characteristics</th>
<th>In-patients</th>
<th>Within 3 months of discharge</th>
<th>Under care of CRHT team</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>( n = 1621 )</td>
<td>( n = 2819 )</td>
<td>( n = 1765 )</td>
</tr>
<tr>
<td>Age, median (range)</td>
<td>44 (15–96)</td>
<td>45 (15–95)</td>
<td>48 (15–95)</td>
</tr>
<tr>
<td>Gender, n (%)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Male</td>
<td>1031 (64)</td>
<td>1800 (64)</td>
<td>1082 (61)</td>
</tr>
<tr>
<td>Female</td>
<td>590 (36)</td>
<td>1019 (36)</td>
<td>683 (39)</td>
</tr>
<tr>
<td>Primary diagnosis, n (%)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Schizophrenia and other delusional disorders</td>
<td>483 (30)</td>
<td>455 (16)</td>
<td>229 (13)</td>
</tr>
<tr>
<td>Bipolar affective disorder</td>
<td>190 (12)</td>
<td>300 (11)</td>
<td>161 (9)</td>
</tr>
<tr>
<td>Depressive illness</td>
<td>602 (37)</td>
<td>990 (35)</td>
<td>845 (48)</td>
</tr>
<tr>
<td>Other diagnosis</td>
<td>346 (21)</td>
<td>1074 (38)</td>
<td>530 (30)</td>
</tr>
<tr>
<td>Method, n (%)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Hanging/strangulation</td>
<td>728 (45)</td>
<td>1175 (42)</td>
<td>813 (46)</td>
</tr>
<tr>
<td>Self-poisoning</td>
<td>143 (9)</td>
<td>644 (23)</td>
<td>357 (20)</td>
</tr>
<tr>
<td>Other methods</td>
<td>750 (46)</td>
<td>1000 (38)</td>
<td>595 (34)</td>
</tr>
<tr>
<td>History of self-harm, n (%)</td>
<td>1227 (76)</td>
<td>2124 (75)</td>
<td>1223 (69)</td>
</tr>
<tr>
<td>History of drug and/or alcohol misuse, n (%)</td>
<td>759 (47)</td>
<td>1492 (53)</td>
<td>743 (42)</td>
</tr>
</tbody>
</table>
CRHT, crisis resolution home treatment.
Figure 3 shows the timing of suicide in psychiatric in-patients in relation to day of admission. Patients who died by suicide were less likely to have been admitted to hospital on a Saturday or Sunday than during the week (IRR = 0.52 (95% CI 0.45–0.60), \( P < 0.001 \)). We obtained similar findings when we restricted the analysis to people who had died within 30 days of admission: IRR weekend v. weekdays = 0.65 (95% CI 0.54–0.78), \( P < 0.001 \); or people who had died within 7 days of admission: IRR weekend v. weekdays = 0.70 (95% CI 0.53–0.94), \( P = 0.02 \).

**Discussion**

**Main findings**

We found clear evidence of a weekend effect for suicide deaths among mental health patient groups, but contrary to our hypothesis, the incidence of suicide was actually 12–15% lower at the weekend. Similarly, when we re-ran our analysis for in-patients only based on the day of admission, the incidence of suicide for
patients admitted on a Saturday or Sunday was much lower than for those admitted during the week. We found no evidence for an August peak in suicide related to the changeover of junior doctors.

Strengths and limitations

Our study involved national data collection with excellent coverage, but our findings need to be interpreted in the context of a number of methodological limitations, the most important of which is the purely descriptive design. We were, of course, unable to investigate causal mechanisms using this approach. We did not adjust for potential case mix differences as our analyses were based on people who had died by suicide rather than an at-risk cohort. It is therefore possible, although we think unlikely, that our findings are the result of patients under mental healthcare at the weekend being at lower risk than patients during the working week. The reduced IRRs could reflect fewer people under mental healthcare at the weekend, but in-patients on weekend leave would be captured in our figures (and in fact the proportion of in-patients who died while on agreed leave at weekends and during the week was similar (56% v. 52%, \( \chi^2 = 1.69, P = 0.19 \)). Of course if many more patients were formally discharged from in-patient wards on a Friday this might partly account for the findings, but this would not apply to

Fig. 2 Patient suicide in England by month of year (2001–2013).

(a) In-patients: August v. other months incidence rate ratio (IRR) = 0.99, 95% CI 0.83–1.18, \( P = 0.93 \). (b) Post-discharge patients (suicide within 3 months of in-patient discharge): August v. other months IRR = 1.01, 95% CI 0.88–1.15, \( P = 0.90 \). (c) Patients under crisis resolution home treatment (CRHT): August v. other months IRR = 1.14, 95% CI 0.96–1.37, \( P = 0.14 \).
post-discharge or community patients or our analysis based on day of admission.

Another potential weakness is that we considered the day of death as our main outcome rather than the day of the actual episode of suicidal behaviour. Self-poisoning is one method of suicide for which there may be a lag between the suicide episode and death. However, when we ran additional analyses, we found no significant difference in the temporal pattern of suicide by different methods (interaction terms for self-poisoning v. other methods were non-significant). It is also possible that some deaths were not discovered until after the weekend and so an inaccurate date of death was recorded on the death certificate. However, this is unlikely to apply to the in-patients and patients under the care of home treatment teams (who should have been seen by staff on a frequent basis) or to post-discharge patients who lived with others (around half the post-discharge sample).

**Interpretation of findings**

How might we explain our findings? If reduced medical staffing does indeed account for the weekend effect in acute medical and surgical specialties, it may be that mental health is relatively protected from this because it is more community focused, more multidisciplinary in nature and perhaps less reliant on on-call medical staff out of hours. It could also be that increased social contact with families and others at weekends helps prevent some suicide deaths at this time. It is also worth noting that some previous studies that reported an elevated mortality related to weekend hospital admission actually found a slightly reduced weekend mortality among people who remained in-patients that is consistent with our study.

The much lower incidence of suicide in people admitted at the weekend in our study is interesting. It could relate to a possible reduced threshold for admission in the absence of high-quality weekend cover in community services, which results in ‘lower-risk’ patients being admitted. A recent paper that suggests shorter admissions for patients admitted to a psychiatric bed at the weekend is consistent with this.

Suicide is a complex phenomenon with a variety of causes, and another explanation for our findings could be that wider societal and environmental factors are more important determinants of suicide than mental health service provision. However, aspects of psychiatric services can be related to suicide, and we have previously shown an association between staffing turnover and suicide rates in UK mental health services. The patients included in this study had high levels of morbidity and need – the majority had significant psychiatric illness and a history of previous suicidal behaviour, and almost half had a history of alcohol or drug misuse. They died in close proximity to care, and we focused on them in the current study because they might be expected to be the groups most vulnerable to changes in care and supervision. We think it is unlikely that the drivers of temporal variation in suicide in these complex clinical groups are identical to the drivers in the general population. Nonetheless, it is possible that any service-related changes in this study were masked by the more general temporal variations in suicide. Previous general population and clinical studies have also found peaks in suicide at the beginning of the working week and in spring. However, when we ran a post hoc analysis based on all suicide deaths in the general population in England (2001–2013) we found that the weekend v. weekday difference was actually slightly smaller than the one we found in the clinical groups – incidence of suicide around 8% lower at the weekend in the general population v. 12–15% lower at the weekend in the patient groups (although these differences were not statistically significant when examined using tests of interaction, P-values ranging from 0.18 to 0.56).

Seven-day working for medical staff is currently a policy priority in the NHS and has a number of potential advantages, such as improving access to care, enhancing continuity of support and reducing morbidity. A key aim of such models is to improve quality and save lives. However, our study does not support the claim that safety is compromised at weekends, at least in mental health services.

*Fig. 3 Suicide in England by day of admission (in-patients 2001–2013).*

Weekend v. weekday incidence rate ratio (IRR) = 0.52, 95% CI 0.45–0.60, P = 0.001.
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